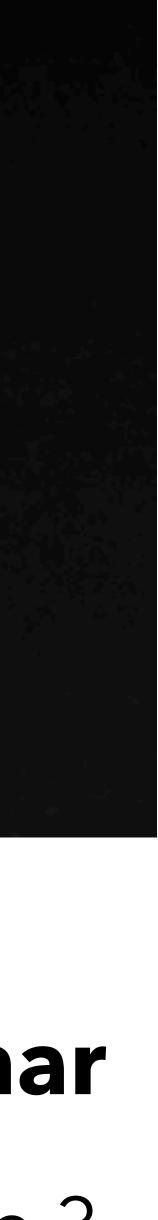


CS4340: Probabilistic Programming Seminar

Lecture 3





- Probabilistic programs are
 - A powerful modelling tool

Recap of the previous lecture

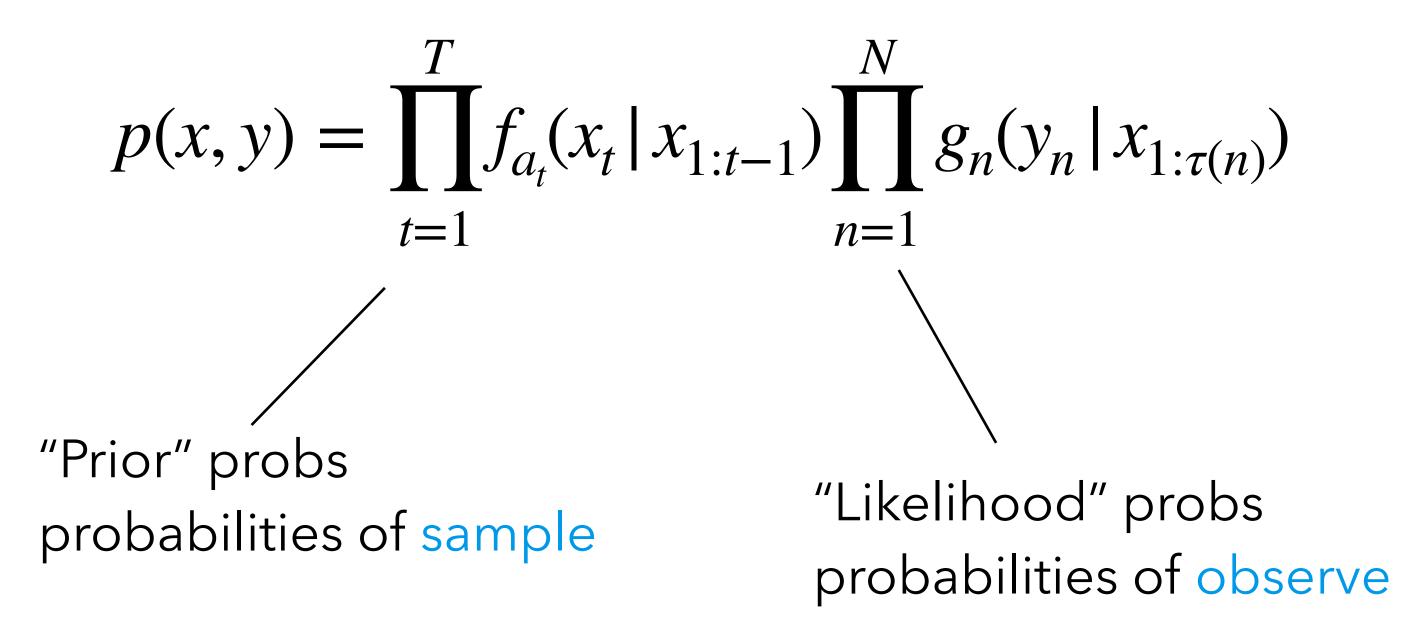
• Programs with two special statements: sample and observe

function probabilisticHelloWorld():

var coin1 = sample(Bernoulli(0.5)) var coin2 = sample(Bernoulli(0.5)) var coin3 = sample(Bernoulli(0.5)) observe(coin2 == 1)

return coin1 + coin2 + coin3 end

Recap of the previous lecture



How do we calculate p(x, y) efficiently?

This lecture

How do we calculate p(x, y) efficiently?

Importance sampling

Particle filtering

This lecture

- Metropolis-Hastings MCMC

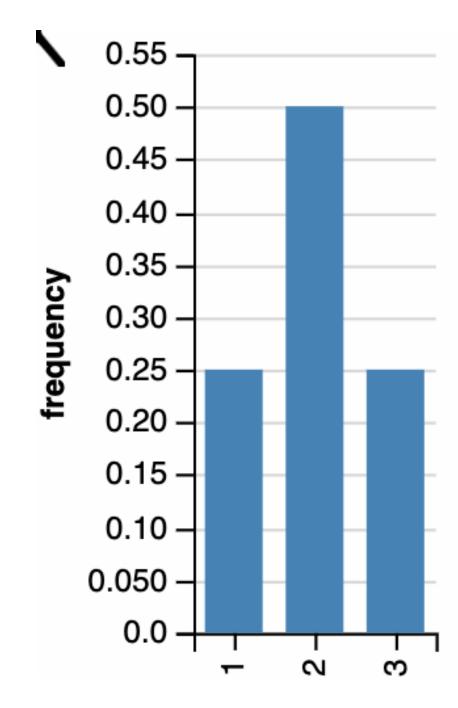
Probabilistic inference

function probabilisticHelloWorld():

var coin1 = sample(Bernoulli(0.5))
var coin2 = sample(Bernoulli(0.5))
var coin3 = sample(Bernoulli(0.5))
observe(coin2 == 1)

return coin1 + coin2 + coin3

end



How do we calculate p(x, y) efficiently?

Importance sampling

Particle filtering

This lecture

- Metropolis-Hastings MCMC

Probabilistic inference Intuition

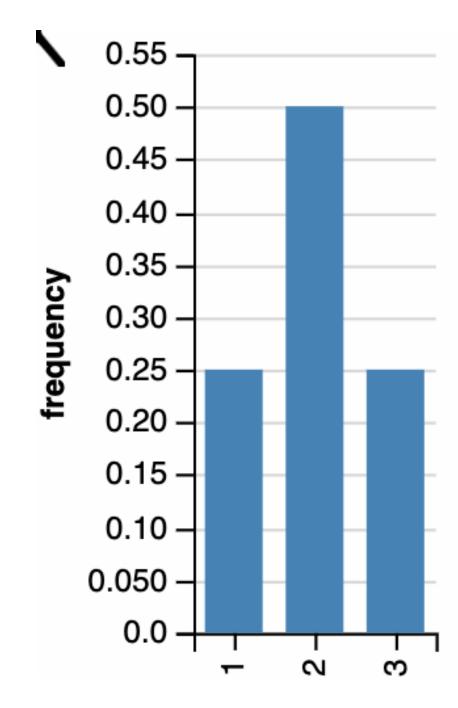
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end

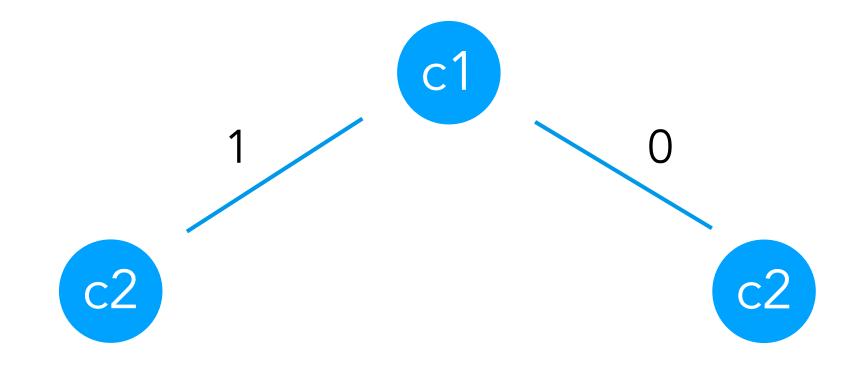


- c1 ~ sample(Bernoulli(0.6))
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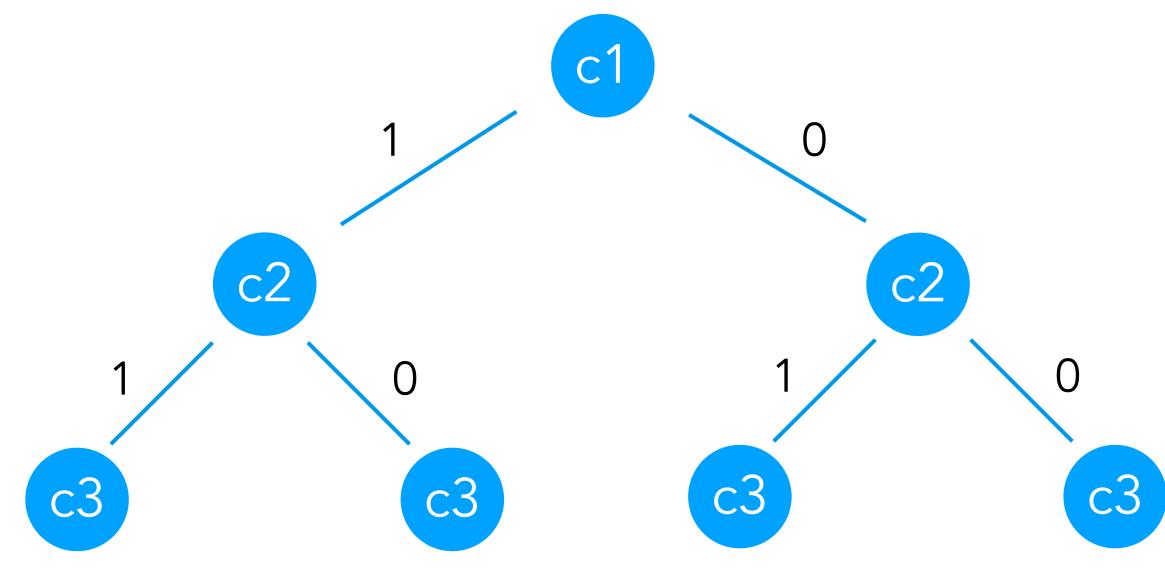
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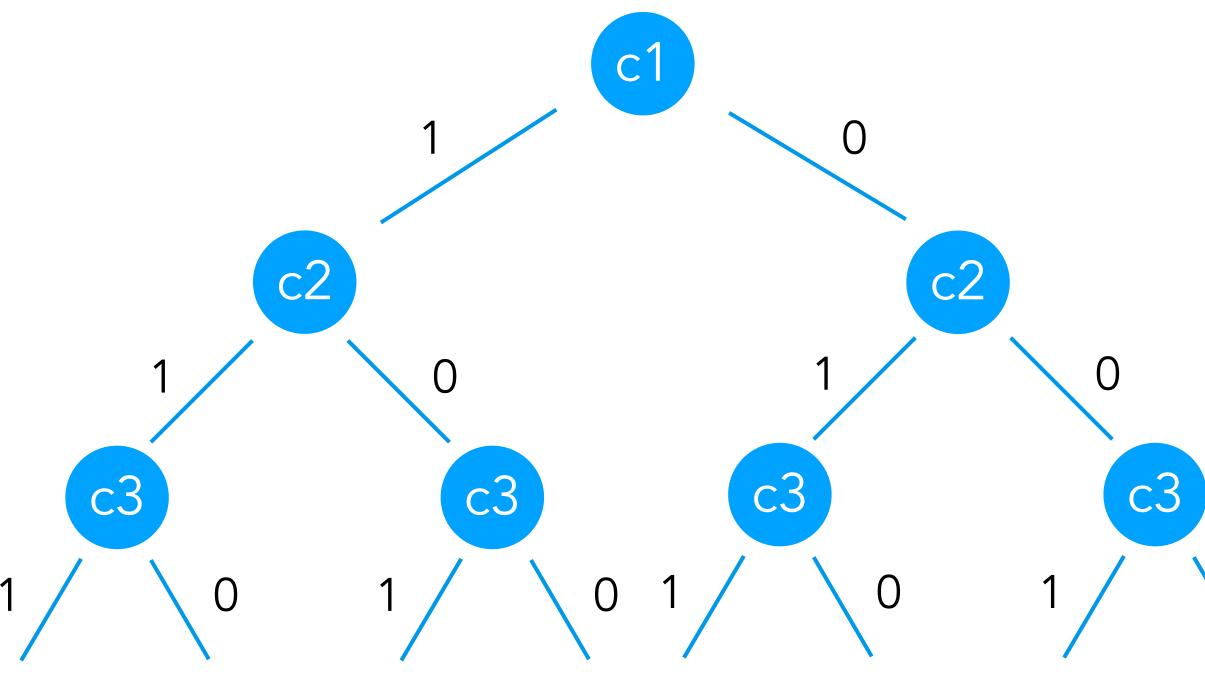


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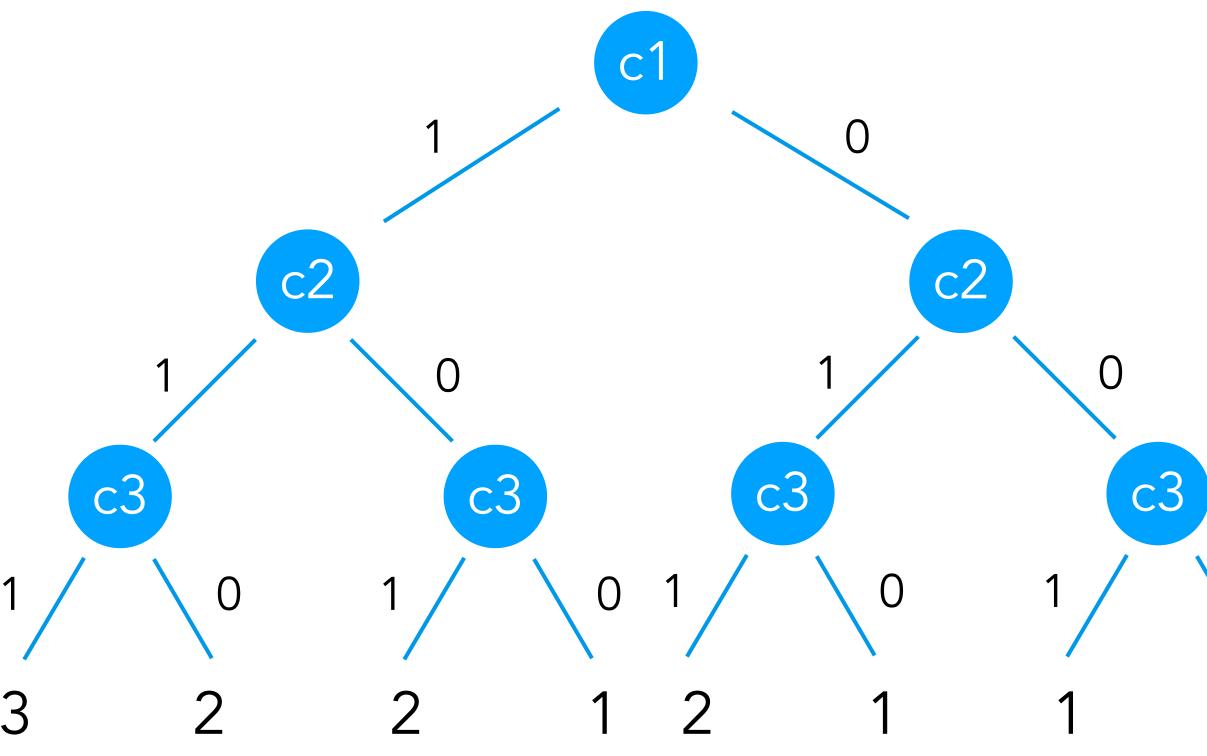


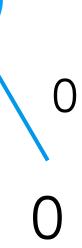
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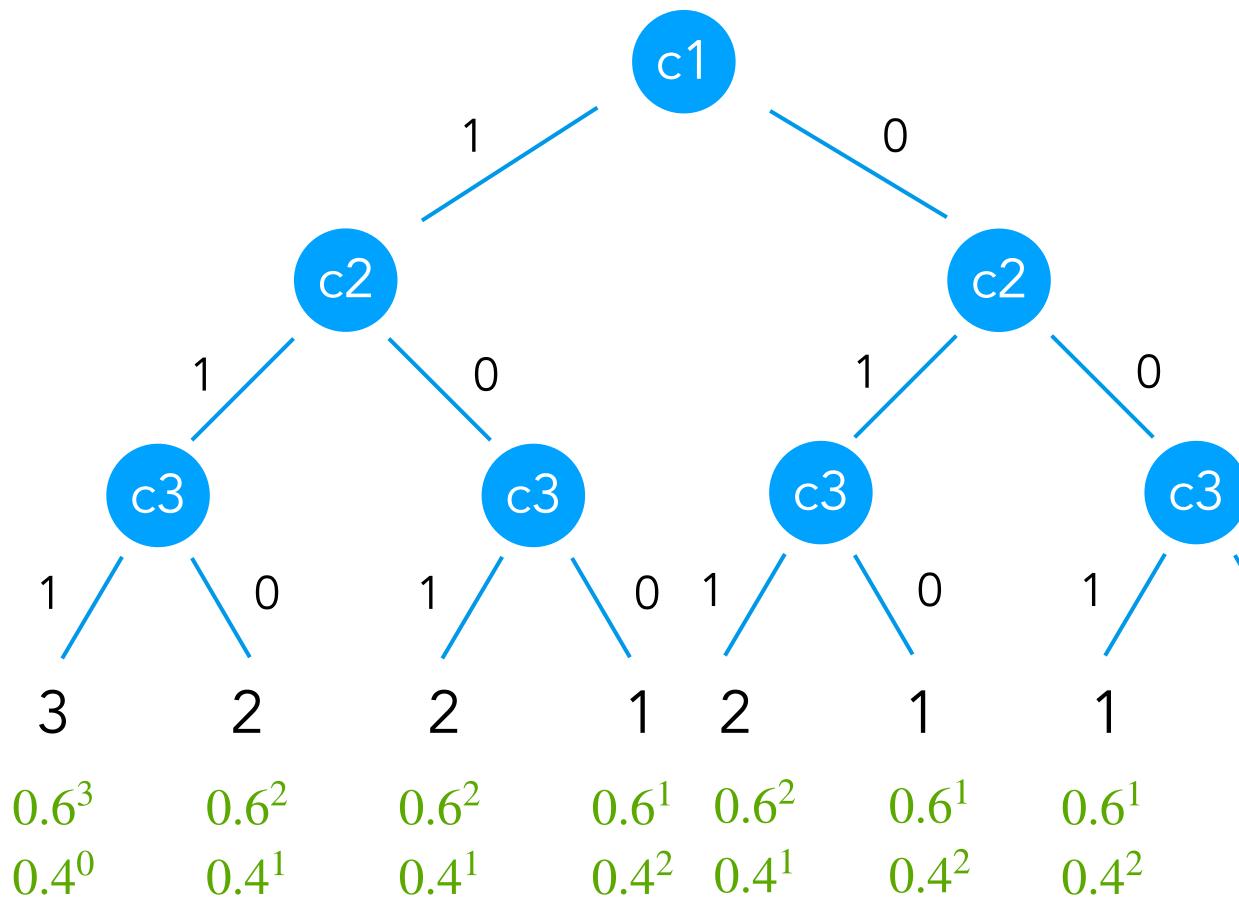


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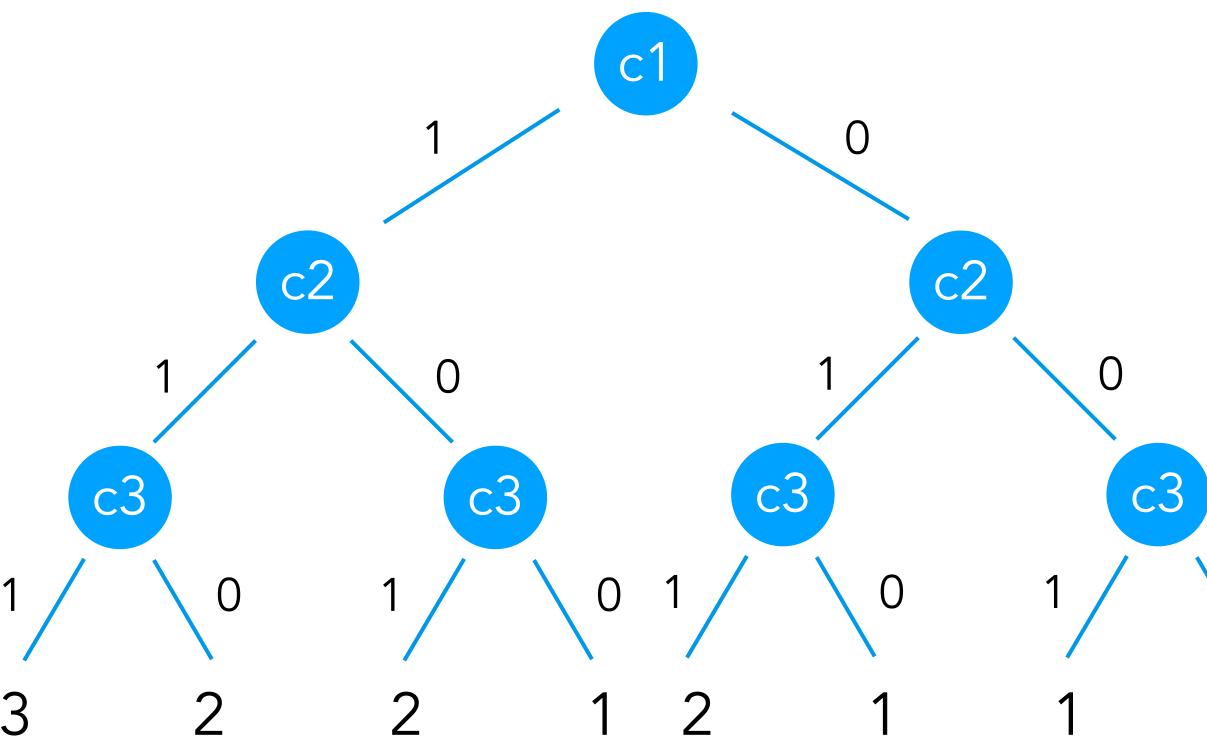
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- c1 ~ sample(Bernoulli(0.6))
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return c1 + c2 + c3

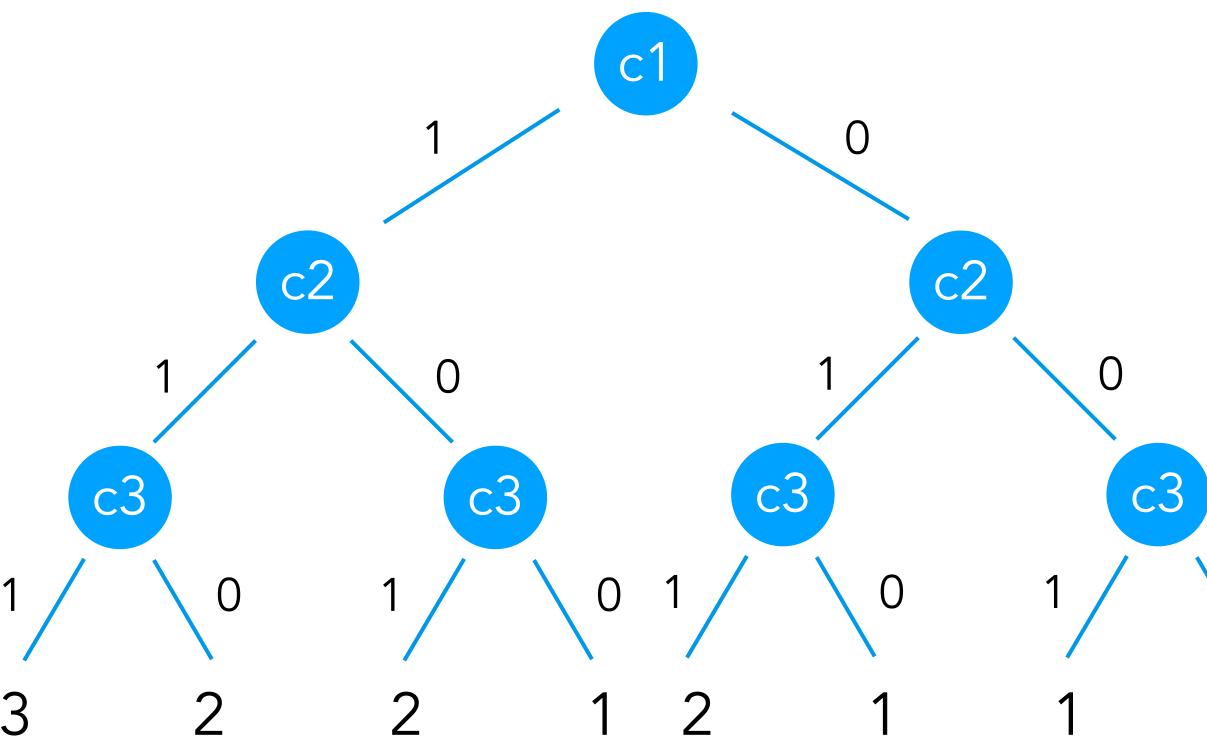


0.216 0.144 0.144 0.096 0.144 0.096 0.096 0.064



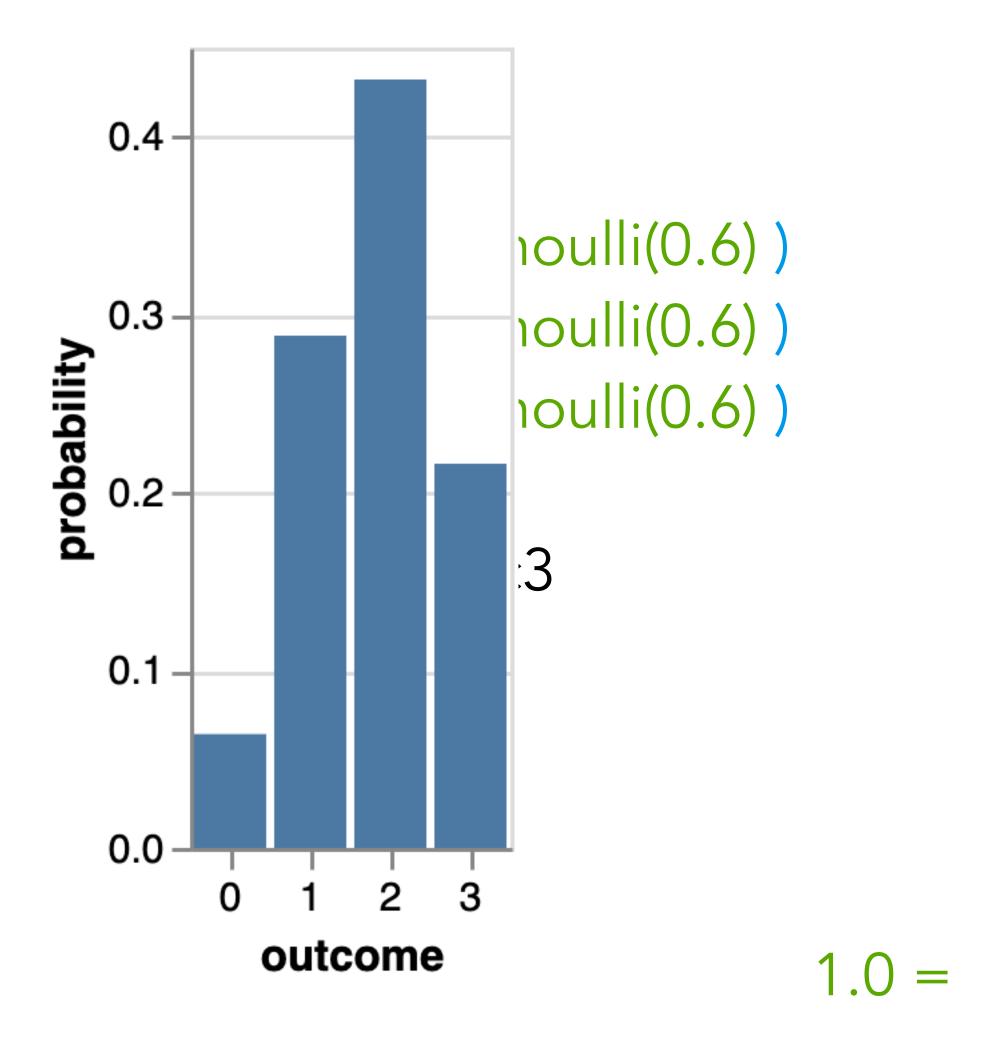
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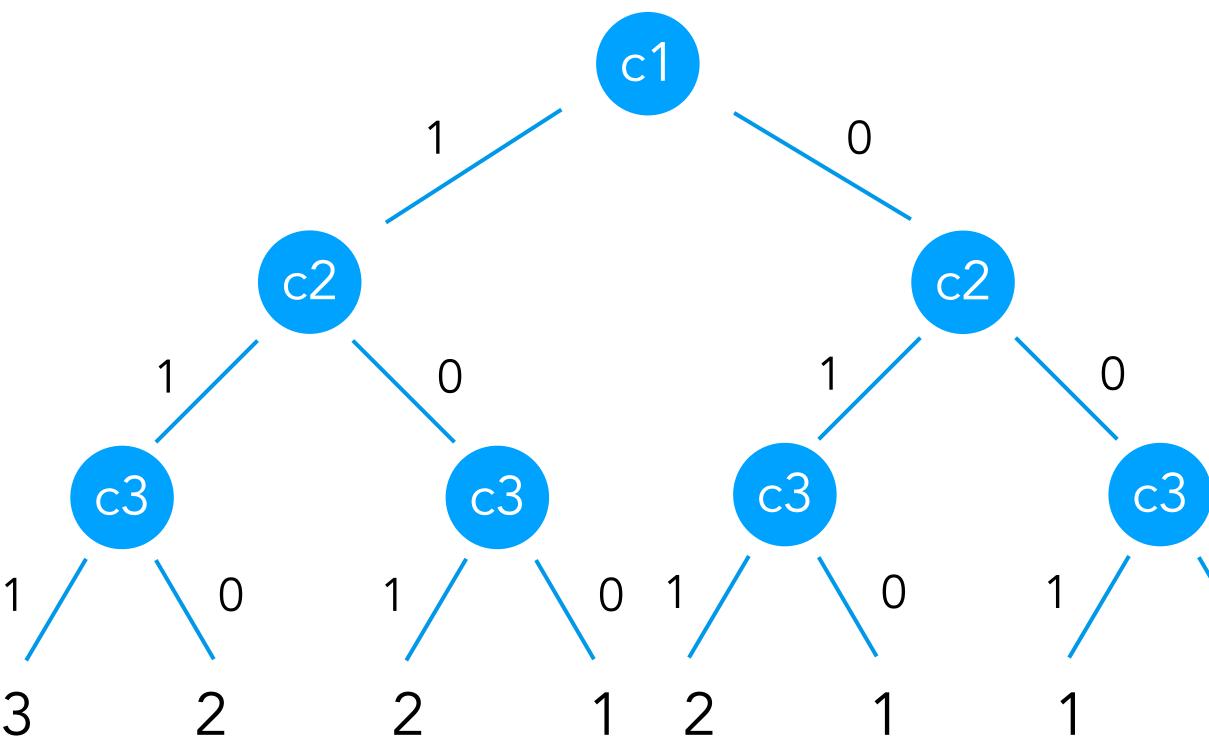
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 $1.0 = 0.216 \quad 0.144 \quad 0.144 \quad 0.096 \quad 0.144 \quad 0.096 \quad 0.096 \quad 0.096 \quad 0.064$





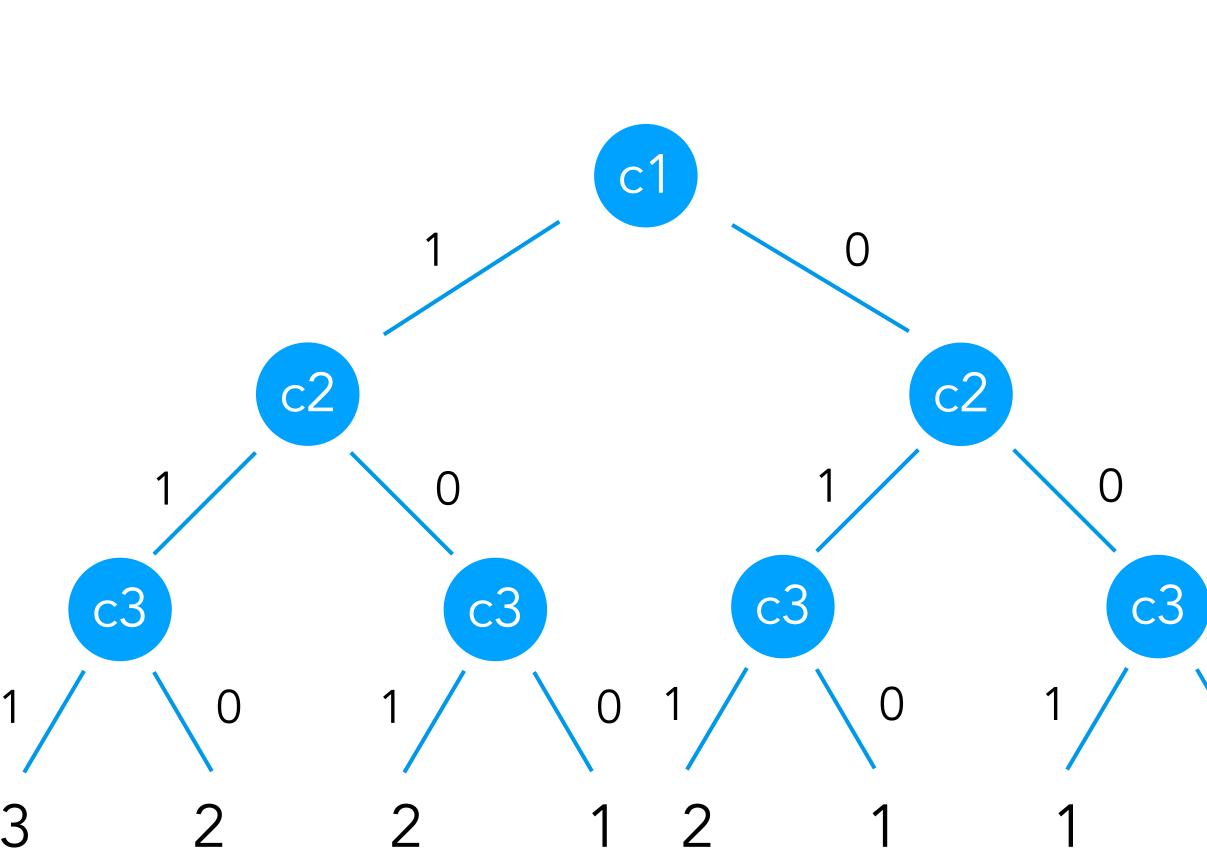


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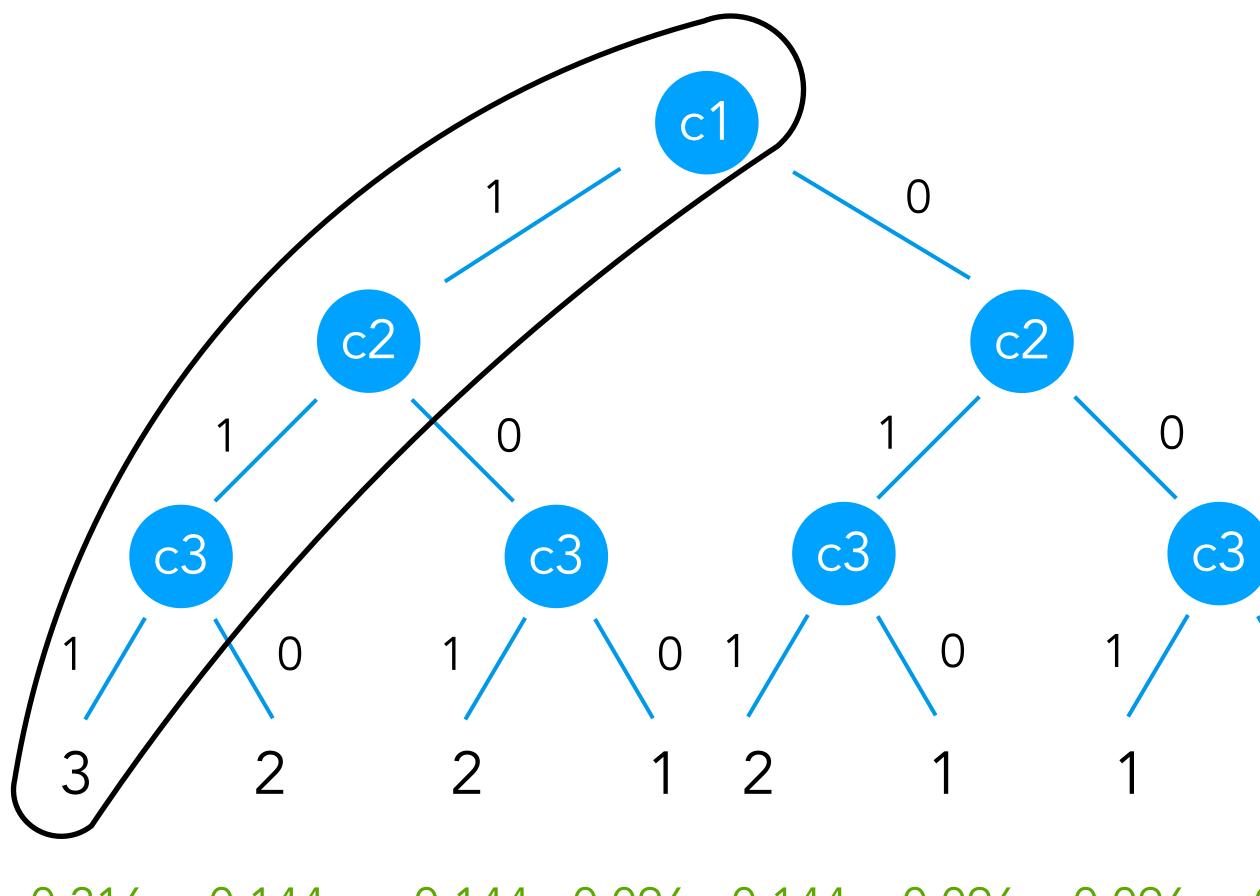
0.144 0.144 0.096 0.144 0.096 0.096 0.064 0.216



Product rule: Probs of random choices multiply

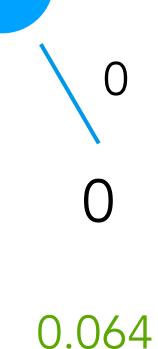
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return c1 + c2 + c3





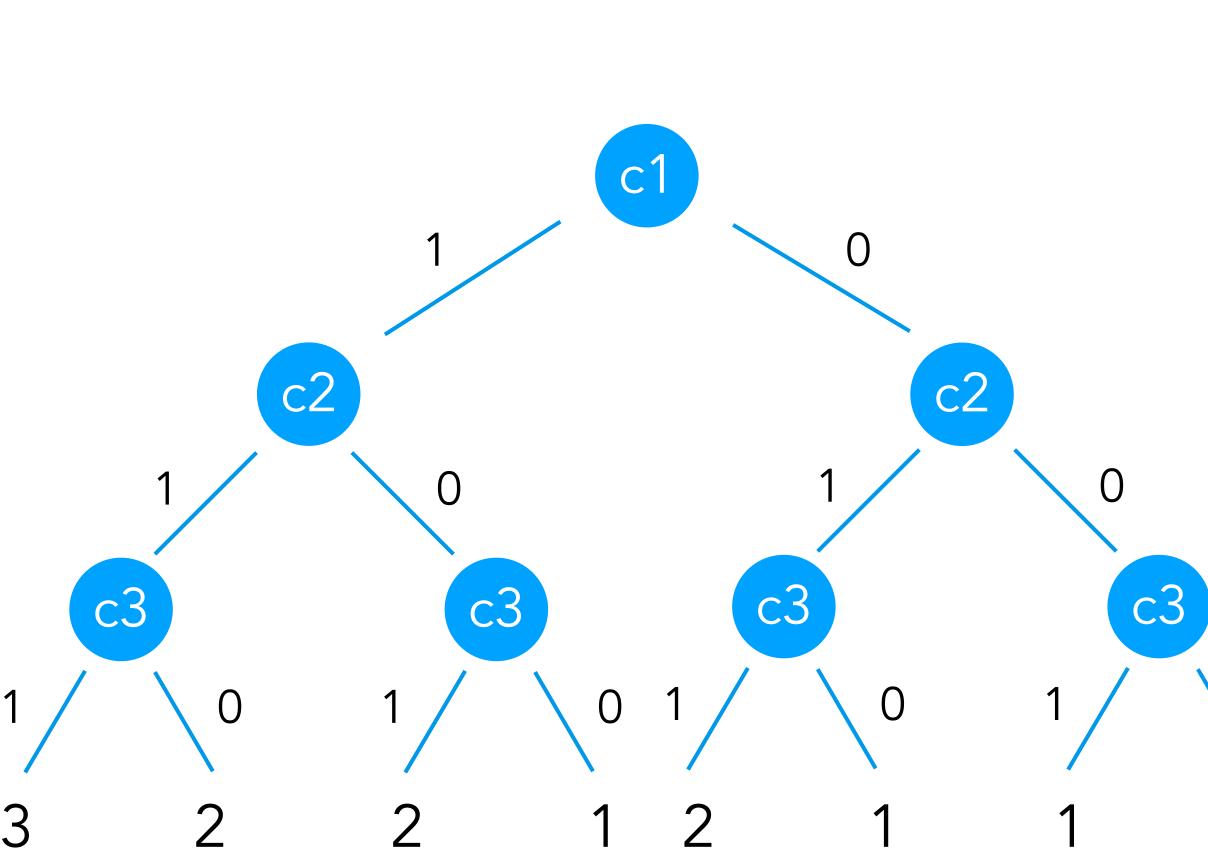
0.216 0.144 0.144 0.096 0.144 0.096 0.096



Product rule: Probs of random choices multiply

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return c1 + c2 + c3



0.144 0.144 0.096 0.144 0.096 0.096 0.216 0.064

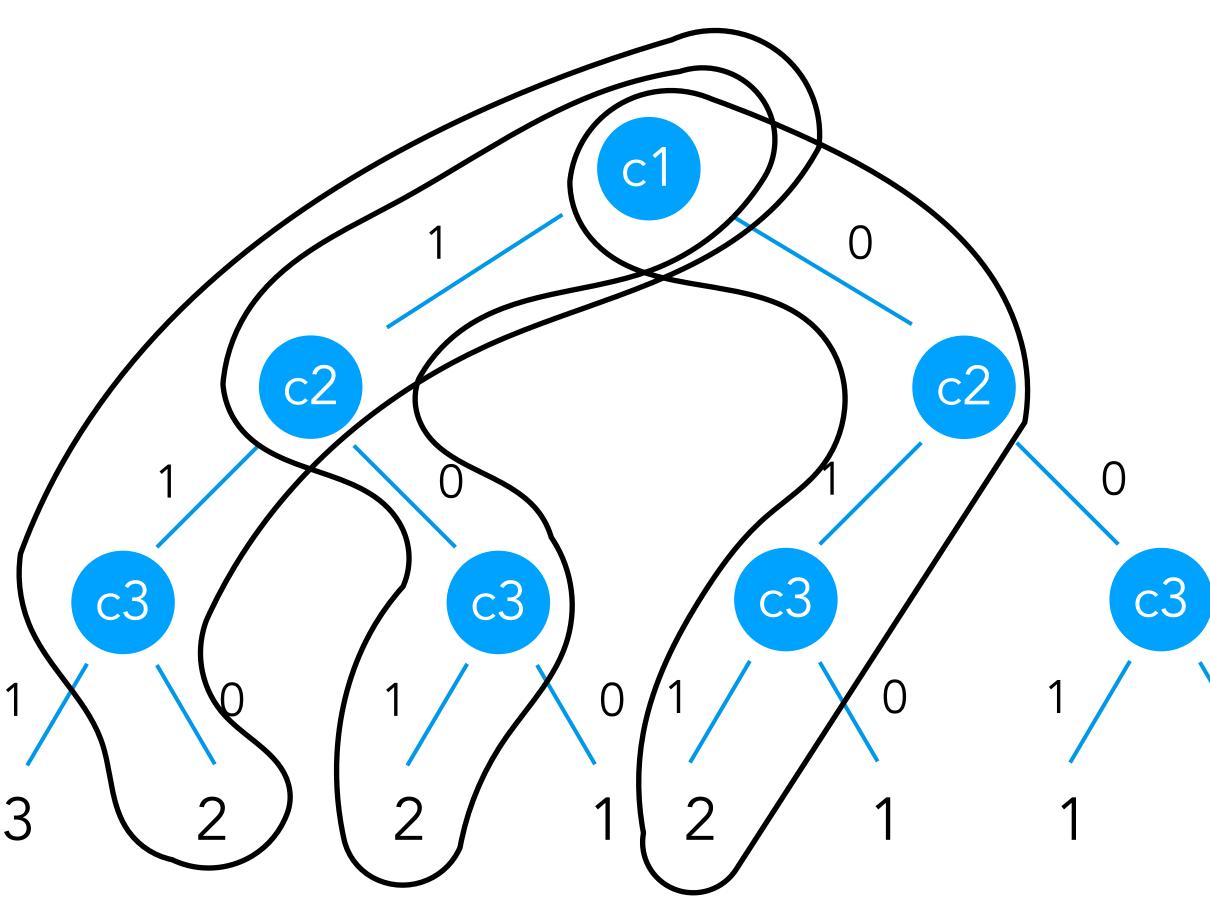


Product rule: Probs of random choices multiply

Sum rule: Probs of alternatives add

- c1 ~ sample(Bernoulli(0.6))
- c2 ~ sample(Bernoulli(0.6))
- c3 ~ sample(Bernoulli(0.6))

return c1 + c2 + c3



0.144 0.096 0.144 0.096 0.096 0.144 0.216



A few exercises

C = [A, B]

return C

A ~ sample(Bernoulli(0.5)) B ~ sample(Bernoulli(0.5))

Probability of C = [true, false]?

A few exercises

C = [A, B]

return C

Probability of C = [true, false]?

A ~ sample(Bernoulli(0.5)) B ~ sample(A ? Bernoulli(0.3) : Bernoulli(0.7))

A few exercises

return C

Probability of C = true ?

C = sample(Bernoulli(0.5)) || sample(Bernoulli(0.5))

- c1 ~ sample(Bernoulli(0.6))
- c2 ~ sample(Bernoulli(0.6))
- c3 ~ sample(Bernoulli(0.6))
- • • •
- c20 ~ sample(Bernoulli(0.6))
- c21 ~ sample(Bernoulli(0.6))

return sum(c1 to c21)

- c1 ~ sample(Bernoulli(0.6))
- c2 ~ sample(Bernoulli(0.6))
- c3 ~ sample(Bernoulli(0.6))
- • • •
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Too many choices to consistently explore

- c1 ~ sample(Bernoulli(0.6))
- c2 ~ sample(Bernoulli(0.6))
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- c20 ~ sample(Bernoulli(0.6))
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Too many choices to consistently explore

We have to approximate: find a representative subset of executions



- c1 ~ sample(Bernoulli(0.6))
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- c3 ~ sample(Bernoulli(0.6))
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Too many choices to consistently explore

We have to approximate: find a representative subset of executions

We approximate with a fixed amount of executions



- c1 ~ sample(Bernoulli(0.6))
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Too many choices to consistently explore

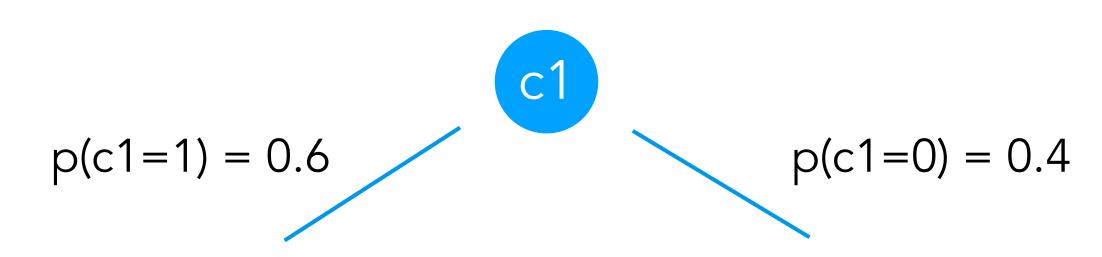
We have to approximate: find a representative subset of executions

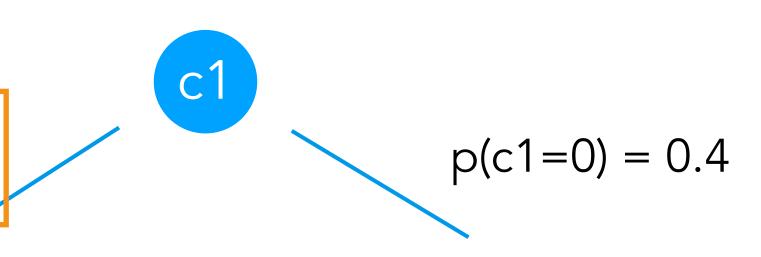
We approximate with a fixed amount of executions

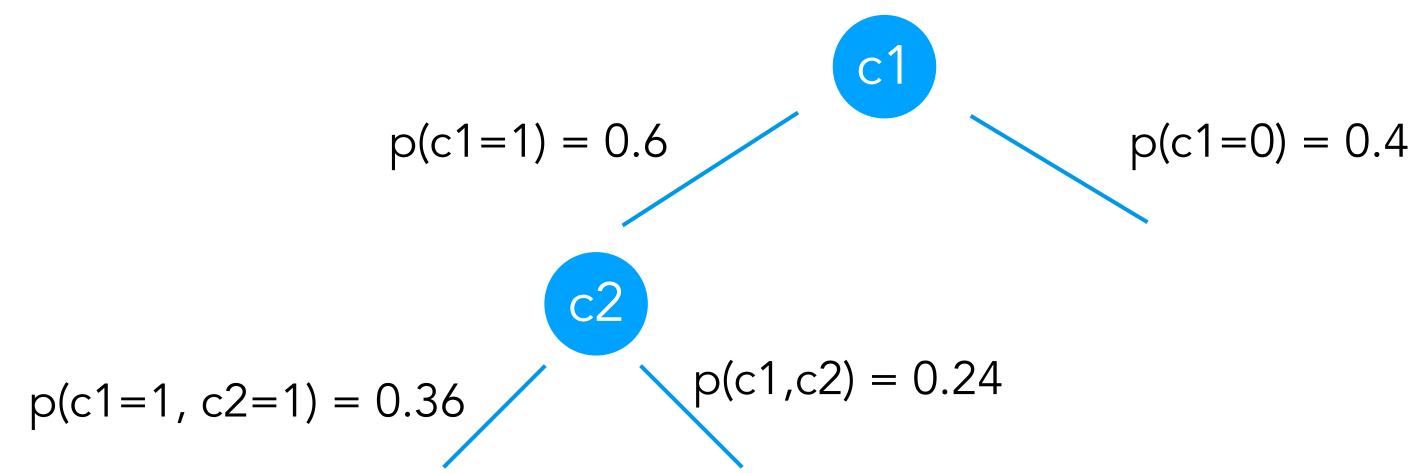
We don't 'care equally about all executions: We care about likely outcome more

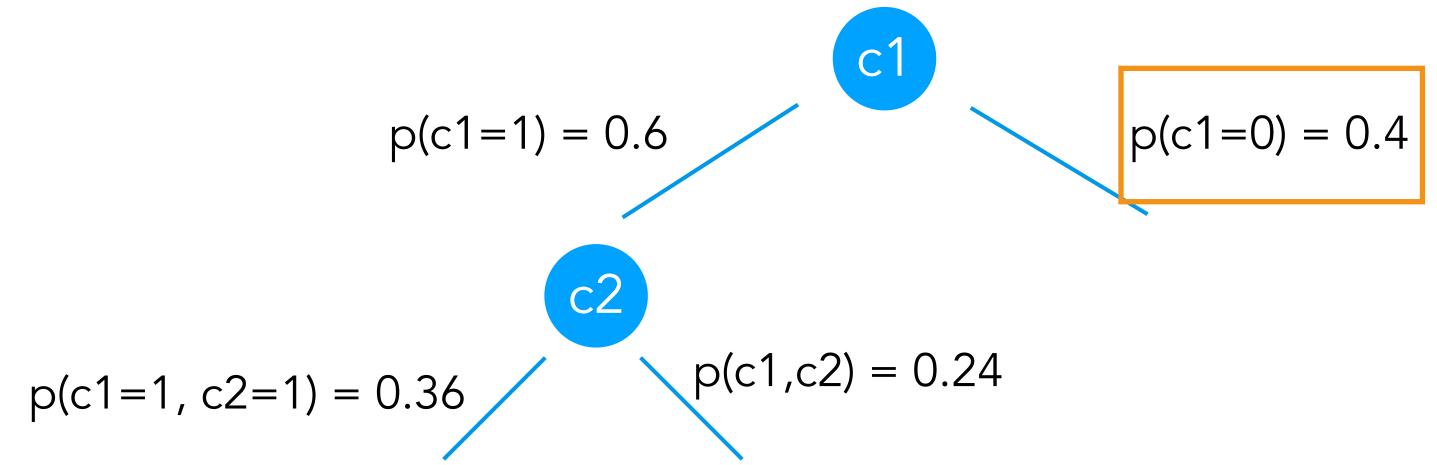


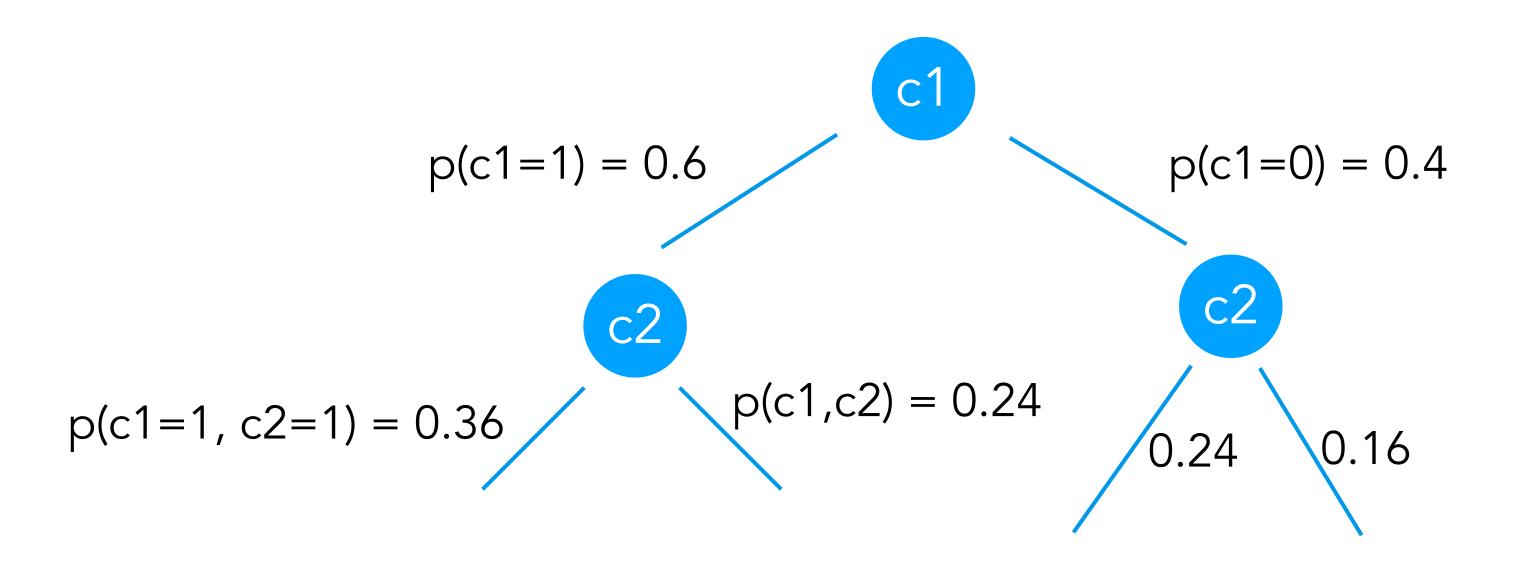


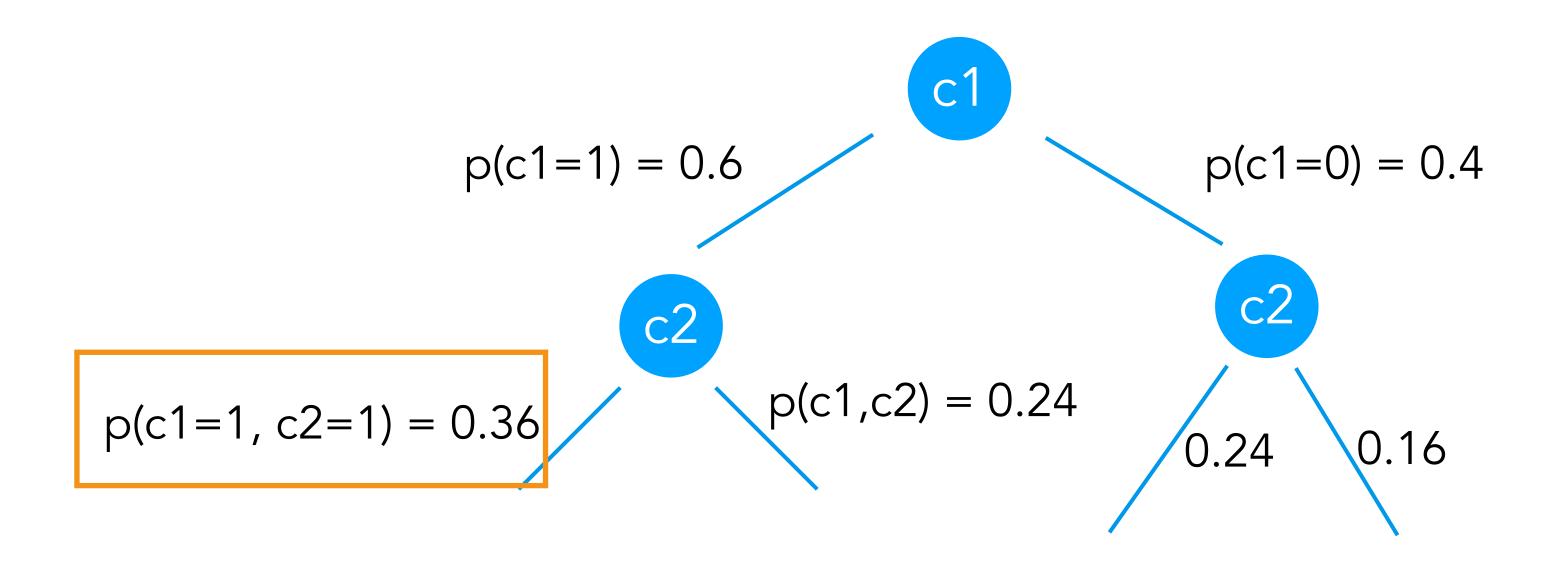


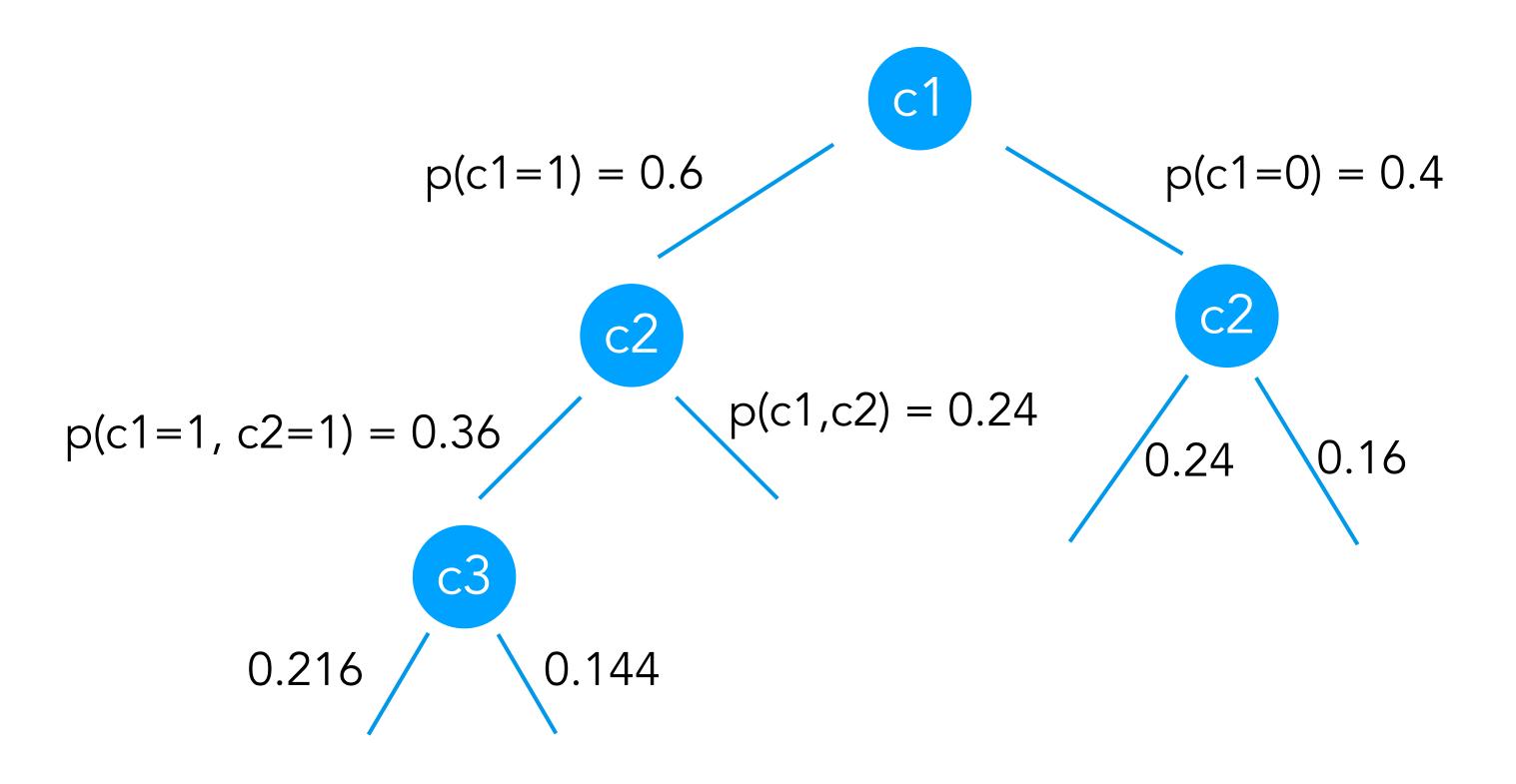




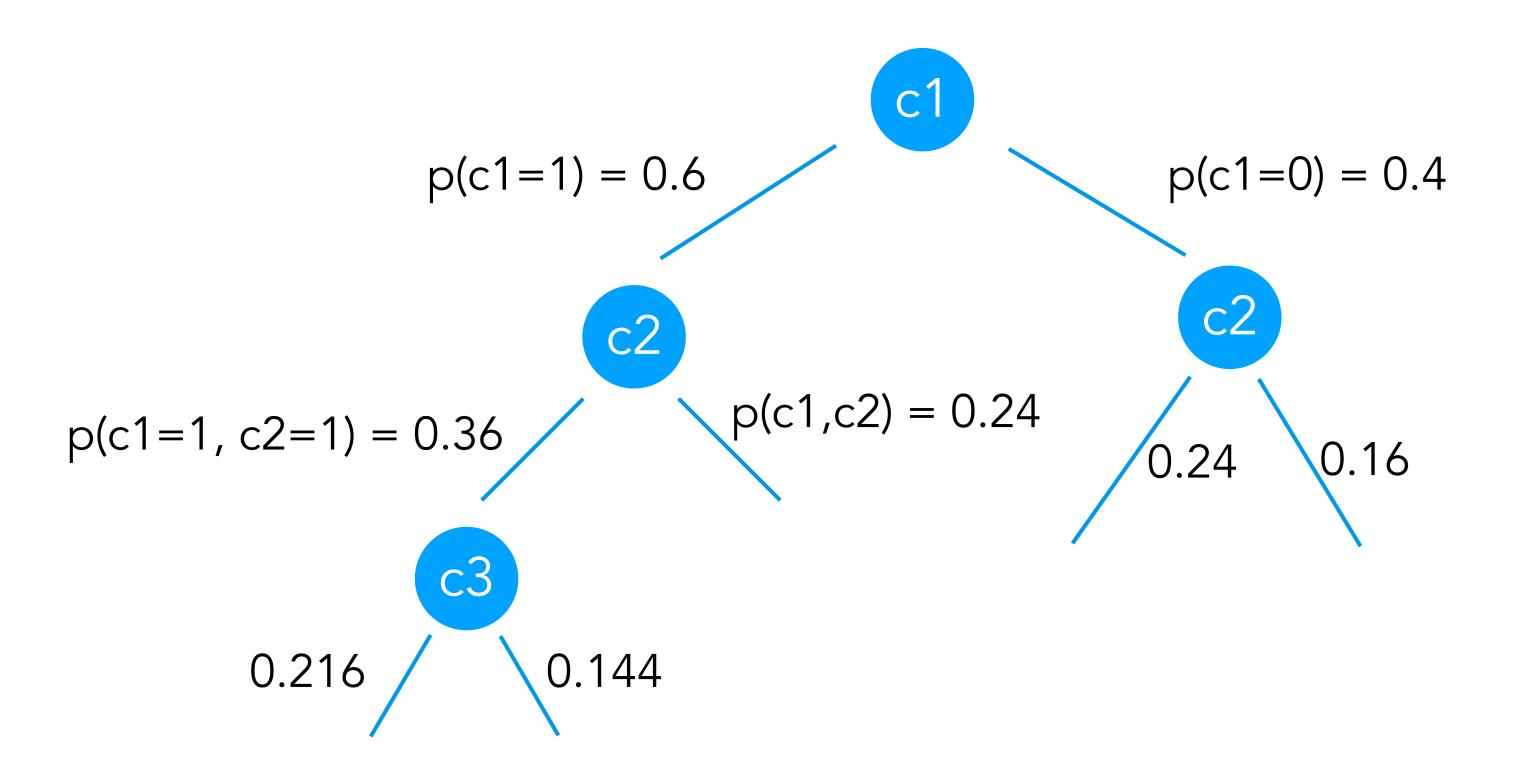








Strategy: order (partial!) executions according to the probabilities of choices



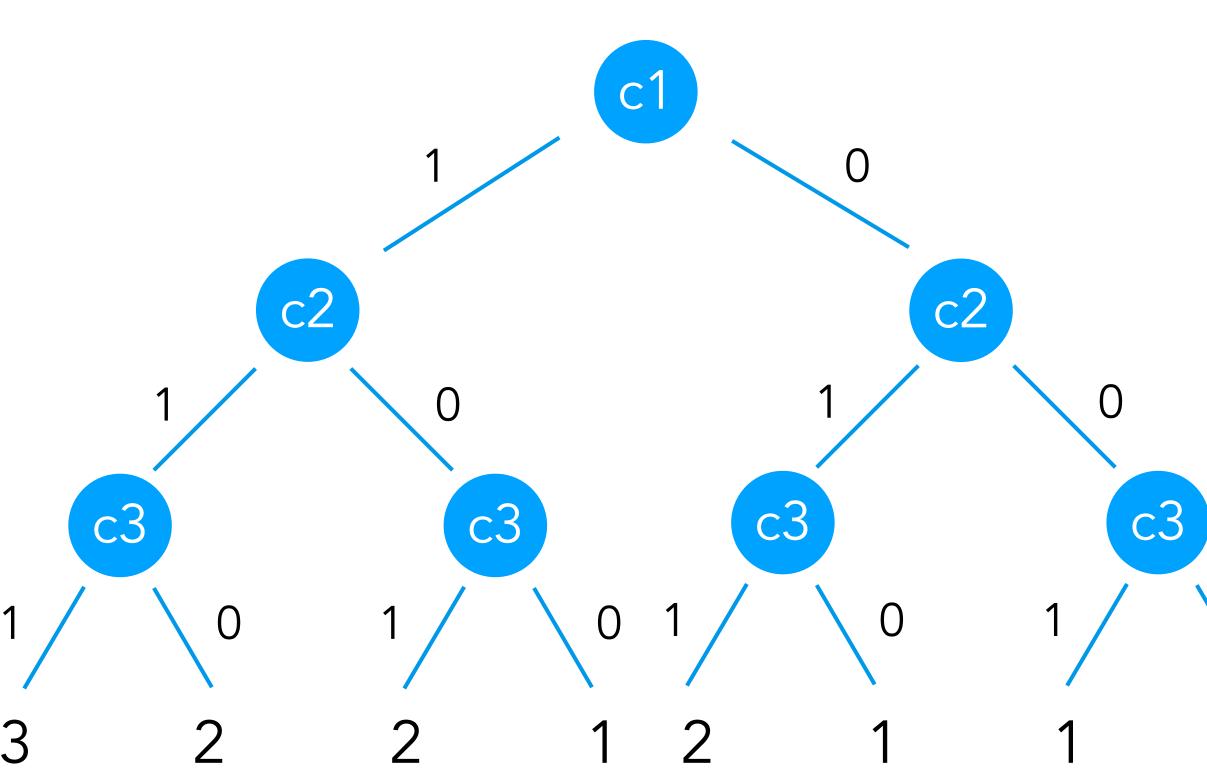
Continue until we collect K executions then normalise

Conditioning:

c1 ~ sample(Bernoulli(0.6)) c2 ~ sample(Bernoulli(0.6)) c3 ~ sample(Bernoulli(0.6))

observe(c2 == 1)

return c1 + c2 + c3



0.216 0.144 0.144 0.096 0.144 0.096 0.096 0.064



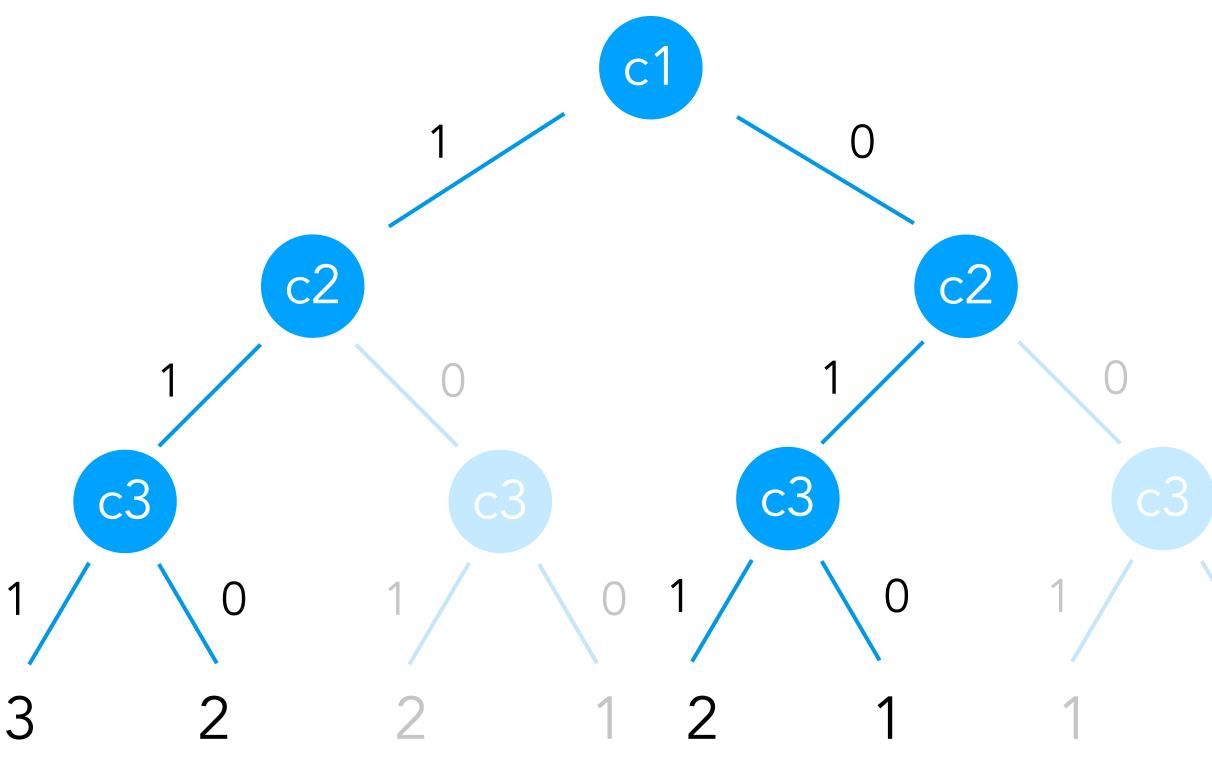
Conditioning: reject violating executions

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0.216 0.144 0.144 0.096 0.144 0.096 0.096 0.064





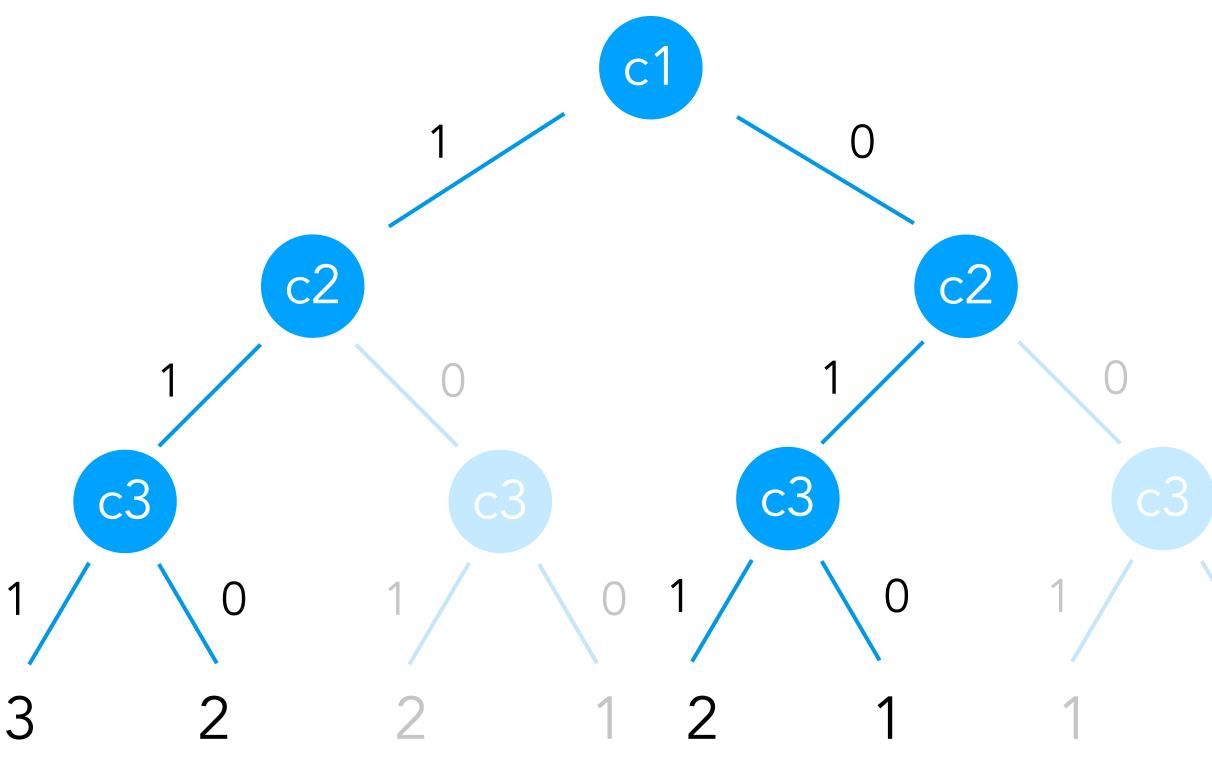
Conditioning: reject violating executions

Valid executions do not sum to 1 anymore

We need to adjust the probabilities according to the Bayes theorem

$$P(A = a | B = b) = \frac{P(A = a, B = b)}{P(B = b)}$$





0.216 0.144 0.144 0.096 0.144 0.096 0.096 0.064

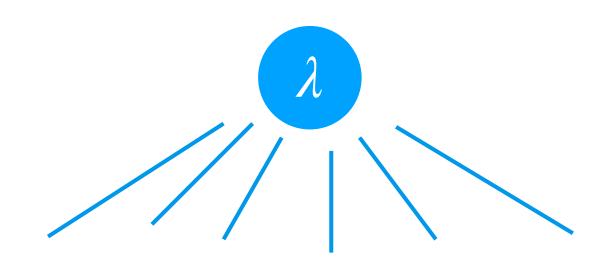




- $\lambda \sim \text{sample}(\text{Normal}(0.5, 1))$
- c1 ~ sample(Bernoulli(λ))
- c2 ~ sample(Bernoulli(λ))
- c3 ~ sample(Bernoulli(λ))

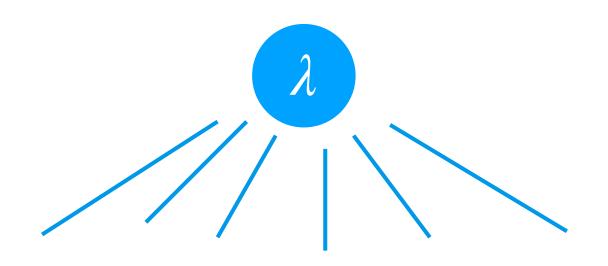


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- c1 ~ sample(Bernoulli(λ))
- c2 ~ sample(Bernoulli(λ))
- c3 ~ sample(Bernoulli(λ))

return c1 + c2 + c3



Infinite number of values

- c1 ~ sample(Bernoulli(λ))
- $c2 \sim sample(Bernoulli(\lambda))$
- $\lambda \sim \text{sample}(\text{Normal}(c1+c2, 0.1))$

observe($\lambda == 2$)

- c1 ~ sample(Bernoulli(λ))
- c2 ~ sample(Bernoulli(λ))
- $\lambda \sim \text{sample}(\text{Normal}(c1+c2, 0.1))$

observe($\lambda == 2$)

- c1 ~ sample(Bernoulli(λ)) c2 ~ sample(Bernoulli(λ))
- $\lambda \sim \text{sample}(\text{Normal}(c1+c2, 0.1))$
- observe(λ , Normal(2, 1))
- return c1 + c2 + c3

- c1 ~ sample(Bernoulli(λ))
- c2 ~ sample(Bernoulli(λ))
- $\lambda \sim \text{sample}(\text{Normal}(c1+c2, 0.1))$

observe($\lambda == 2$)

return c1 + c2 + c3

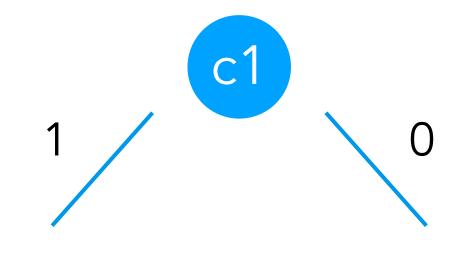
Probability that λ value is an observation from the distribution Normal(2,1)

- c1 ~ sample(Bernoulli(λ)) c2 ~ sample(Bernoulli(λ))
- $\lambda \sim \text{sample}(\text{Normal}(c1+c2, 0.1))$

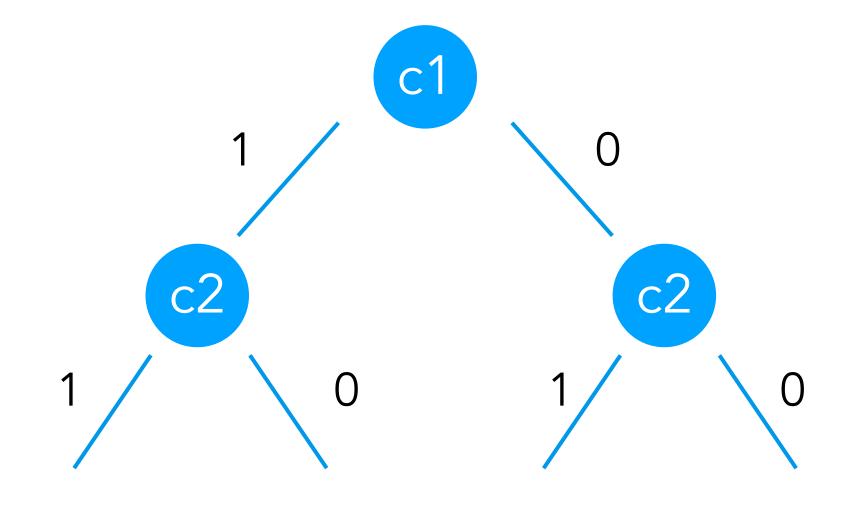
- observe(λ , Normal(2, 1))

- c1 ~ sample(Bernoulli(0.6))
- c2 ~ sample(Bernoulli(0.4))
- $\lambda \sim \text{sample}(\text{Normal}(c1+c2, 0.1))$
- observe(λ , Normal(2, 1))
- return c1 + c2 + c3

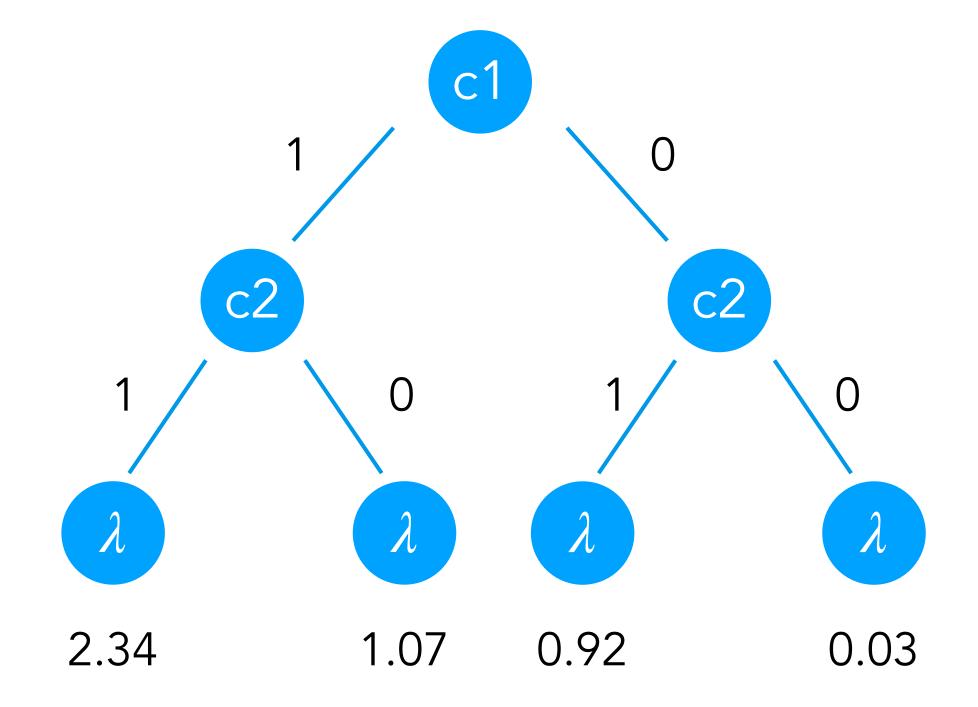
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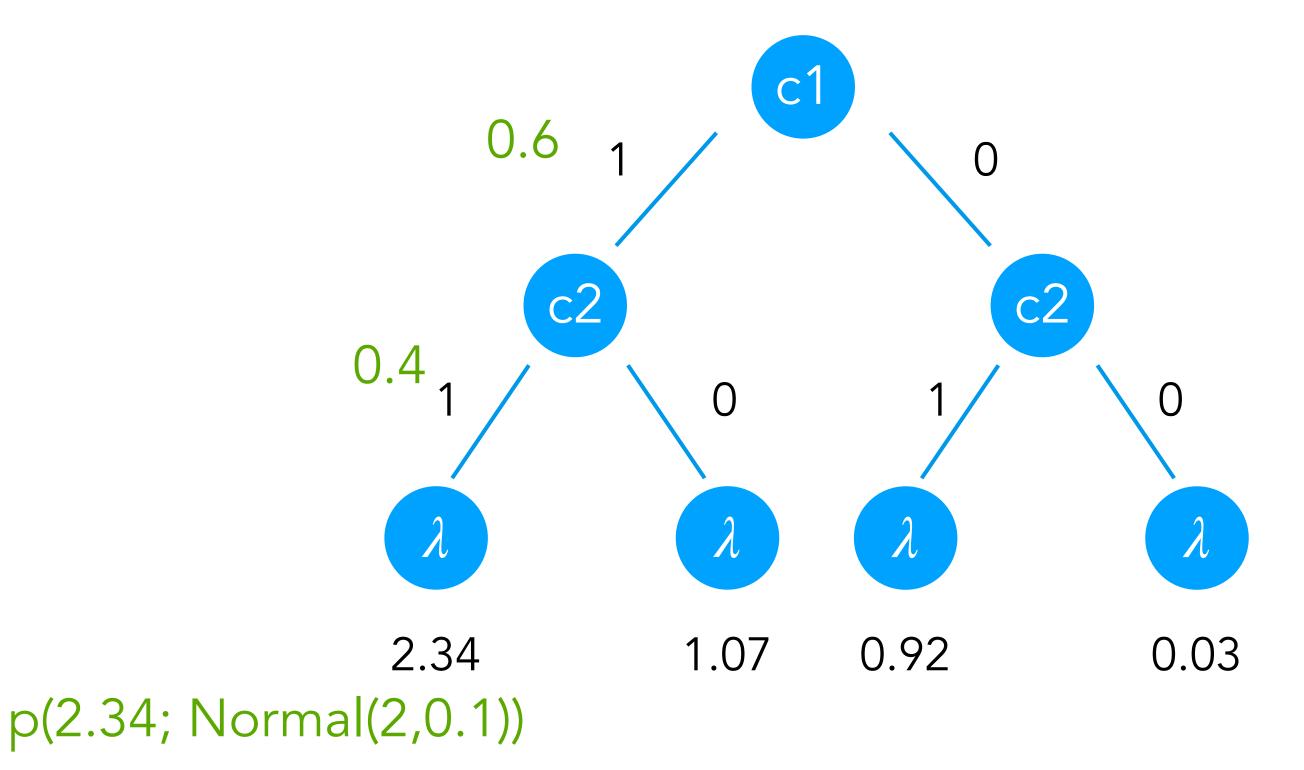
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- $\lambda \sim \text{sample}(\text{Normal}(c1+c2, 0.1))$
- observe(λ, Normal(2, 1))
- return c1 + c2 + c3 p(2.34)



Desiderata for general inference techniques

- General inference technique: doesn't care what is in the program • All programming constructs (loops, conditions, ...) • All distributions (continuous and discrete) • Finite and infinite distribution traces

Probabilistic inference Grand tour

- $\lambda \sim \text{sample}(\text{Normal}(0.5, 1))$
- c1 ~ sample(Bernoulli(λ))
- $c2 \sim sample(Bernoulli(\lambda))$
- $c3 \sim sample(Bernoulli(\lambda))$

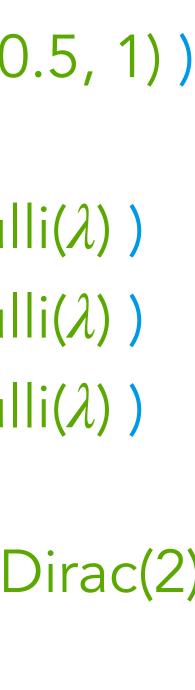
observe(c1+c2+c3, Dirac(2))

return λ

- $\lambda \sim \text{sample(Normal(0.5, 1))}$
- c1 ~ sample(Bernoulli(λ))
- $c2 \sim sample(Bernoulli(\lambda))$
- $c3 \sim sample(Bernoulli(\lambda))$

observe(c1+c2+c3, Dirac(2))

return λ

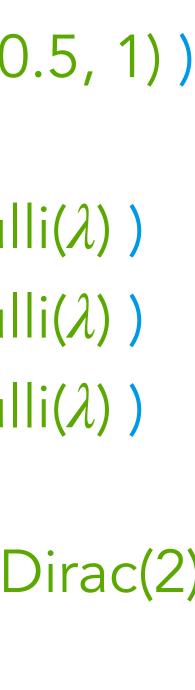


Probabilistic choice

- $\lambda \sim sample(Normal(0.5, 1))$
- c1 ~ sample(Bernoulli(λ))
- $c2 \sim sample(Bernoulli(\lambda))$
- c3 ~ sample(Bernoulli(λ))

observe(c1+c2+c3, Dirac(2))

return λ



Probabilistic choice

 $\lambda \sim sample(Normal(0.5, 1))$

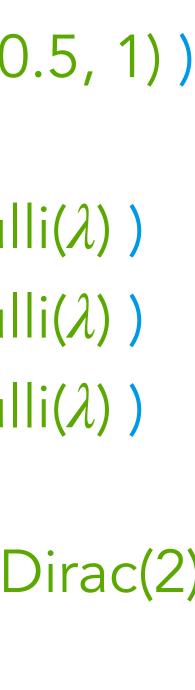
c1 ~ sample(Bernoulli(λ)) c2 ~ sample(Bernoulli(λ))

c3 ~ sample(Bernoulli(λ))

>observe(c1+c2+c3, Dirac(2))

return λ

Observations



Probabilistic choice

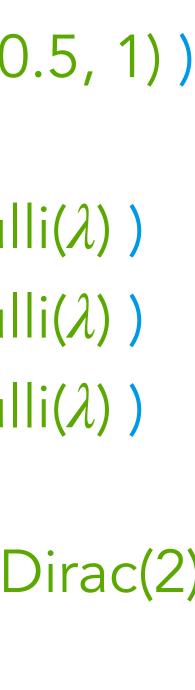
 $\lambda \sim sample(Normal(0.5, 1))$

Observations

c1 ~ sample(Bernoulli(λ)) c2 ~ sample(Bernoulli(λ))

c3 ~ sample(Bernoulli(λ))

>observe(c1+c2+c3, Dirac(2))



Probabilistic choice

 $\lambda \sim sample(Normal(0.5, 1))$

Observations

c1 ~ sample(Bernoulli(λ)) $c2 \sim sample(Bernoulli(\lambda))$

 $c3 \sim sample(Bernoulli(\lambda))$

observe(c1+c2+c3, Dirac(2))

return λ Outcome

Trace: a state of all probabilistic choices in a program

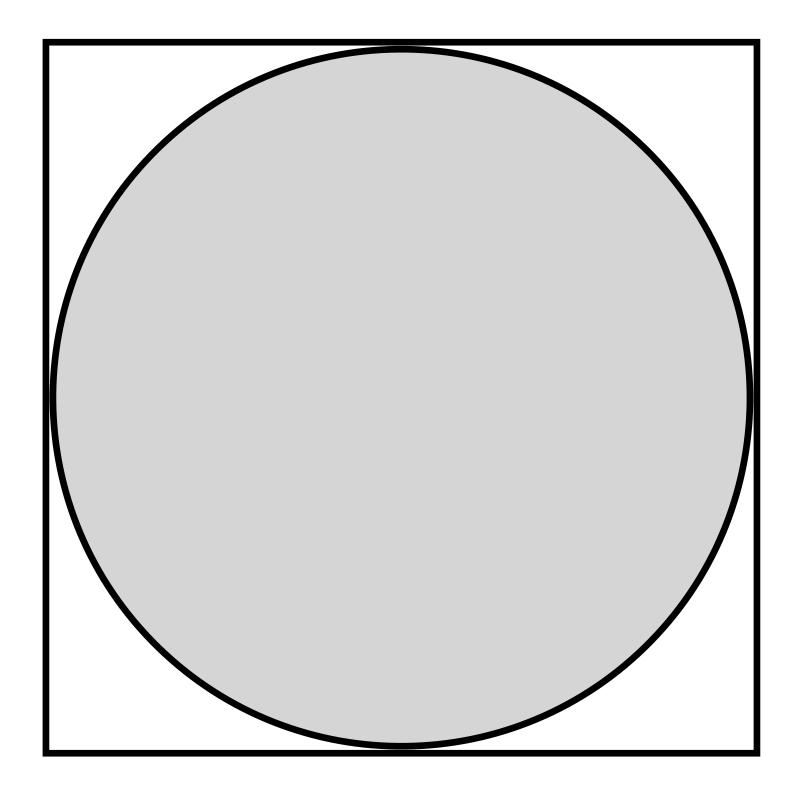
 $\lambda: 0.43$ c1:0 c2:1 c3:0

Randomly simulate a process P(desired outcome) = E[desired outcome] simulations with desired outcome all simulations

Randomly simulate a process

P(desired outcome) = E[desired outcome]

simulations with desired outcome

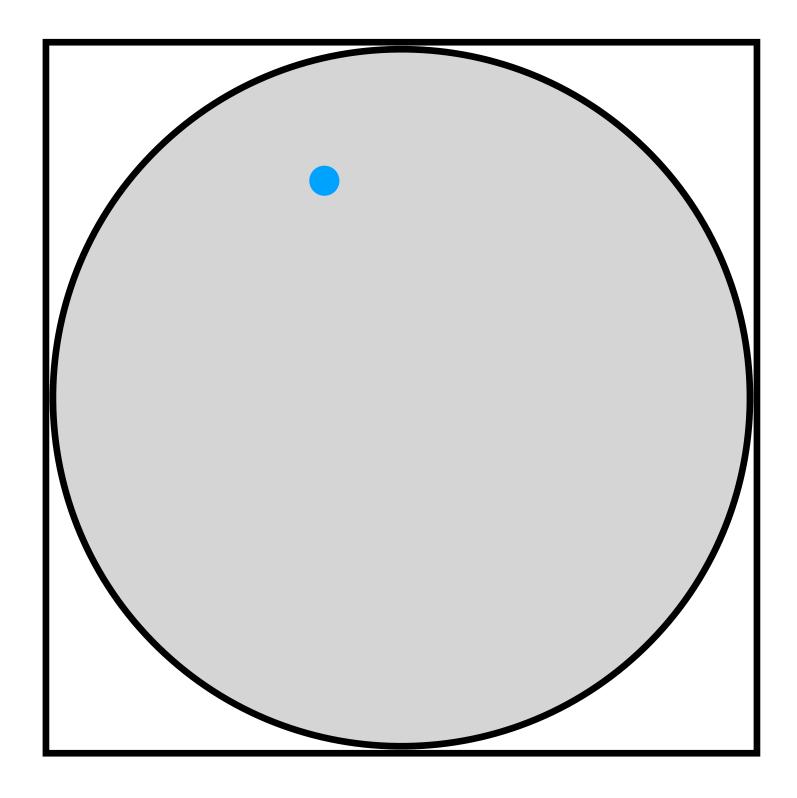


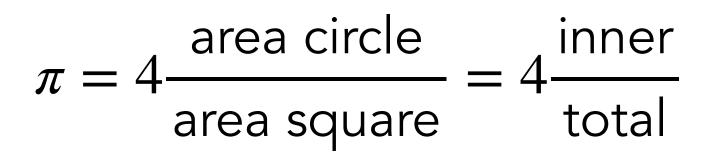
$$\pi = 4 \frac{\text{area circle}}{\text{area square}} = 4 \frac{\text{inner}}{\text{total}}$$

Randomly simulate a process

P(desired outcome) = E[desired outcome]

simulations with desired outcome

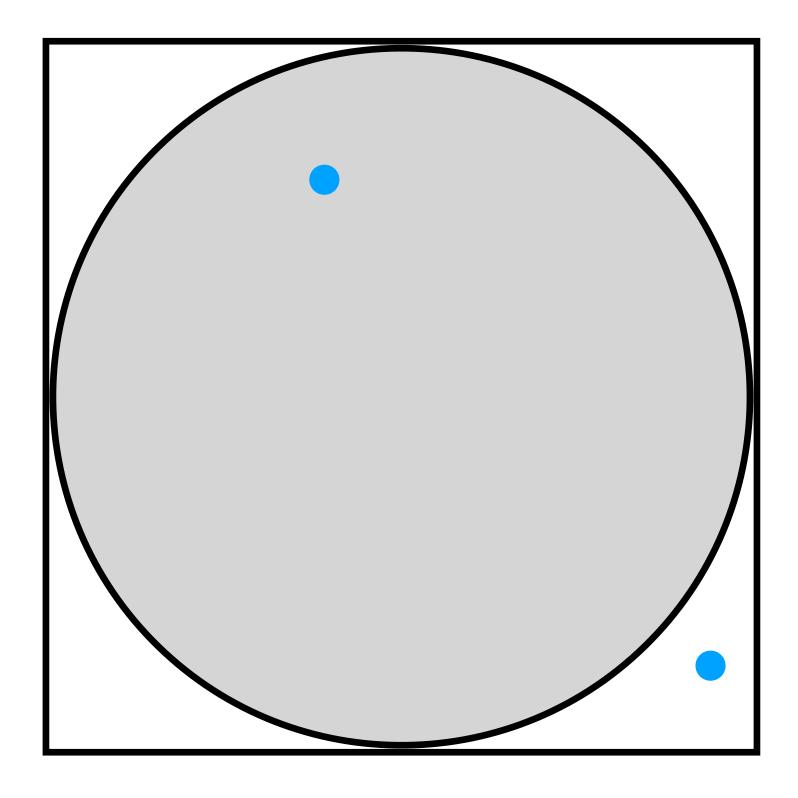




Randomly simulate a process

P(desired outcome) = E[desired outcome]

simulations with desired outcome

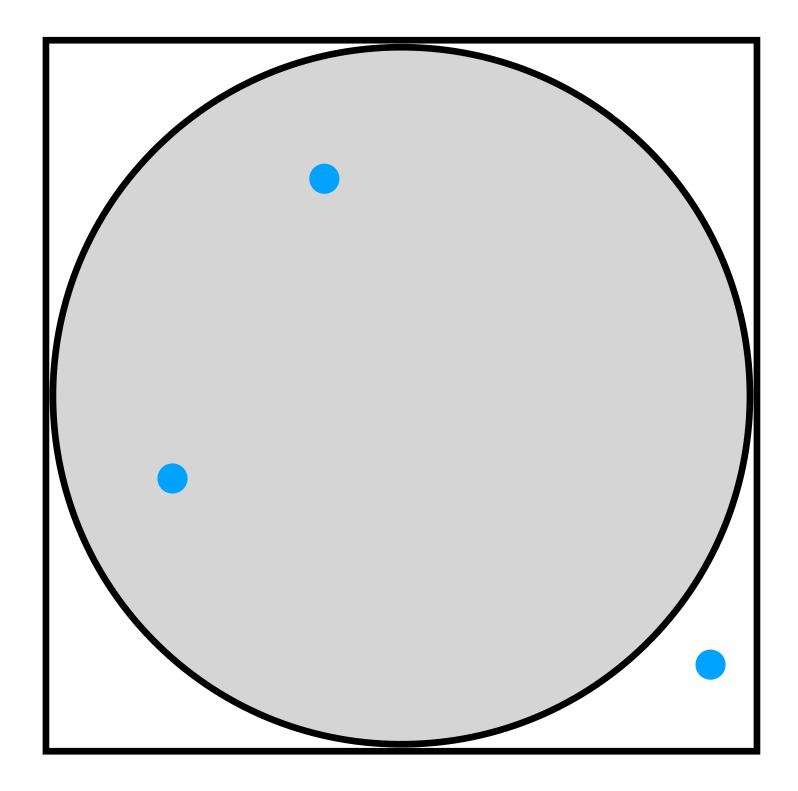


$$\pi = 4 \frac{\text{area circle}}{\text{area square}} = 4 \frac{\text{inner}}{\text{total}}$$

Randomly simulate a process

P(desired outcome) = E[desired outcome]

simulations with desired outcome

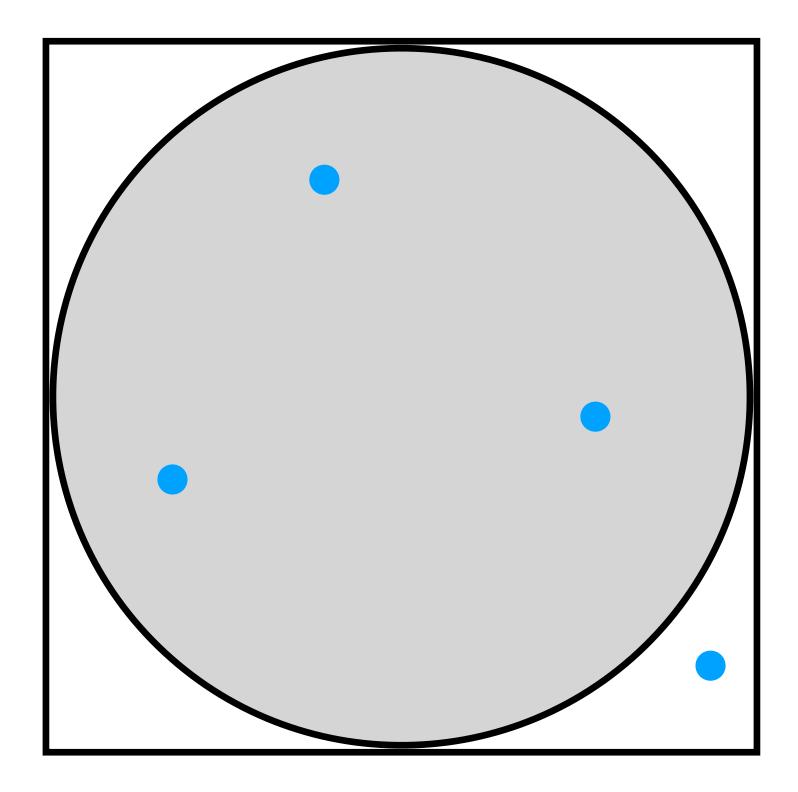


$$\pi = 4 \frac{\text{area circle}}{\text{area square}} = 4 \frac{\text{inner}}{\text{total}}$$

Randomly simulate a process

P(desired outcome) = E[desired outcome]

simulations with desired outcome

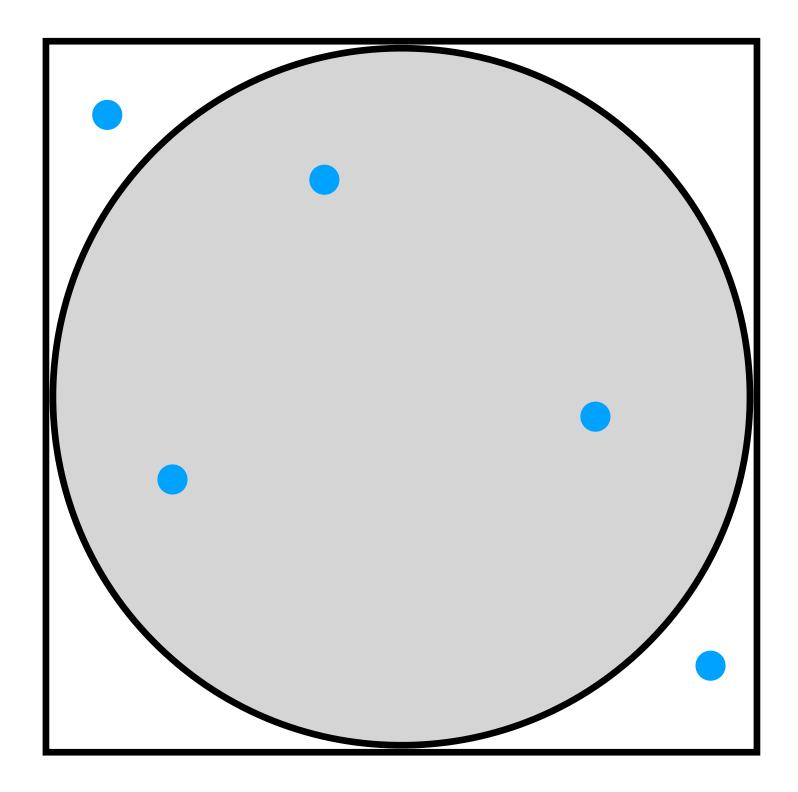


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Randomly simulate a process

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simulations with desired outcome

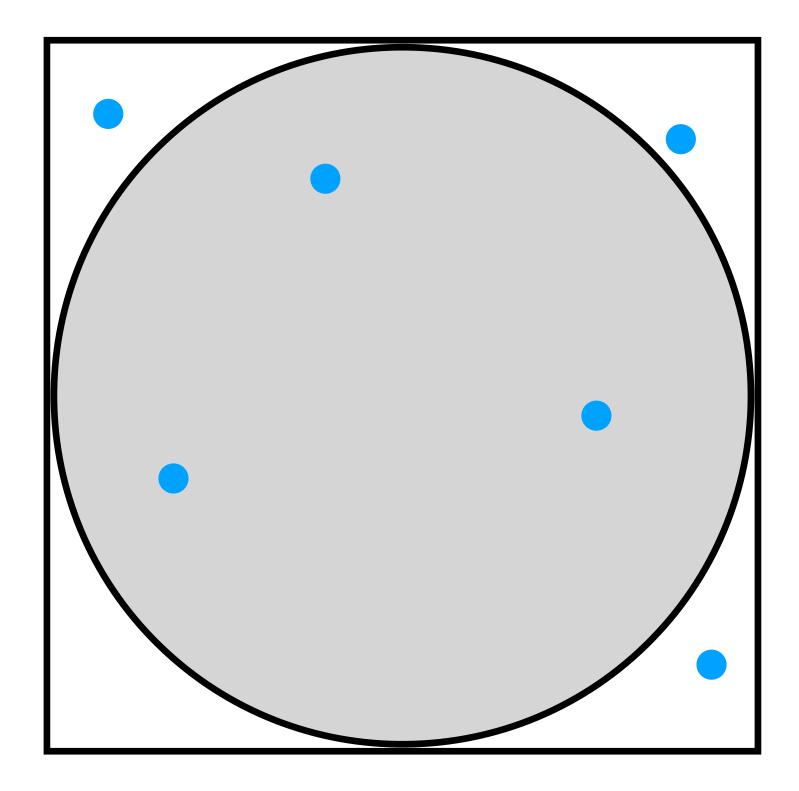


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Randomly simulate a process

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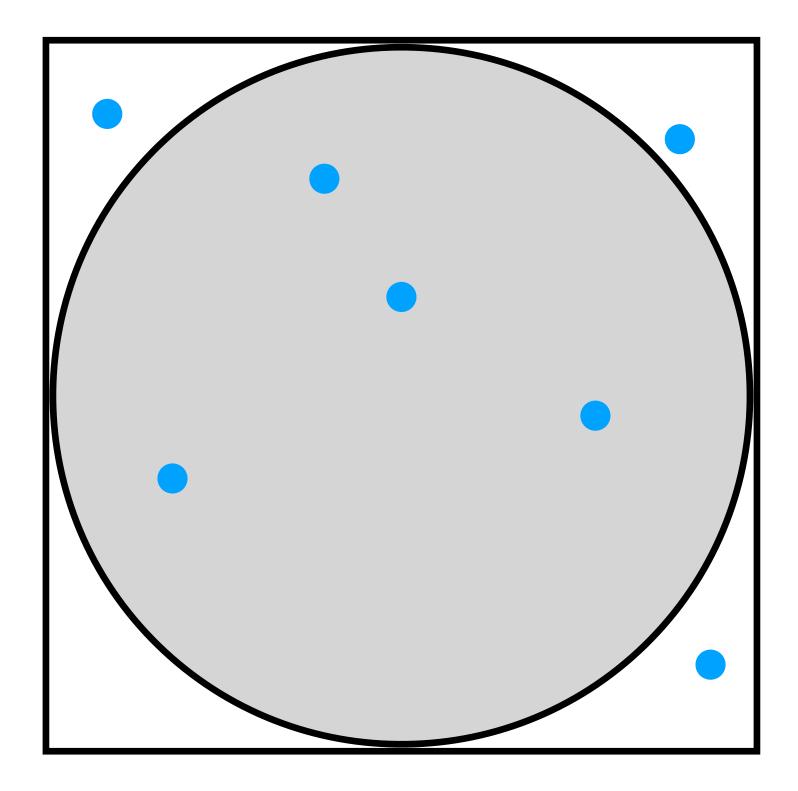


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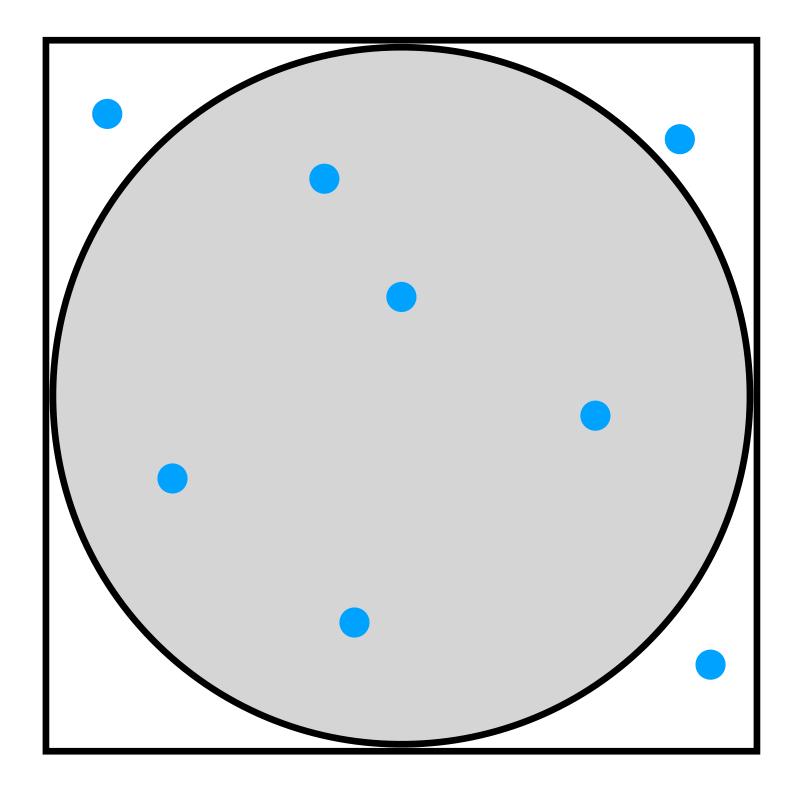


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P(desired outcome) = E[desired outcome]

simulations with desired outcome

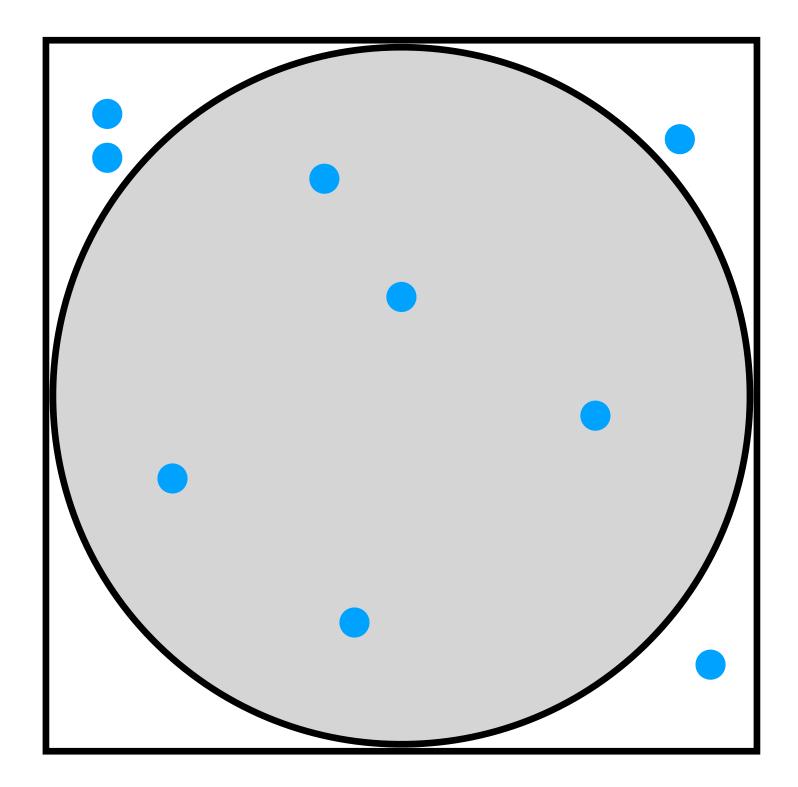


$$\pi = 4 \frac{\text{area circle}}{\text{area square}} = 4 \frac{\text{inner}}{\text{total}}$$

Randomly simulate a process

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simulations with desired outcome

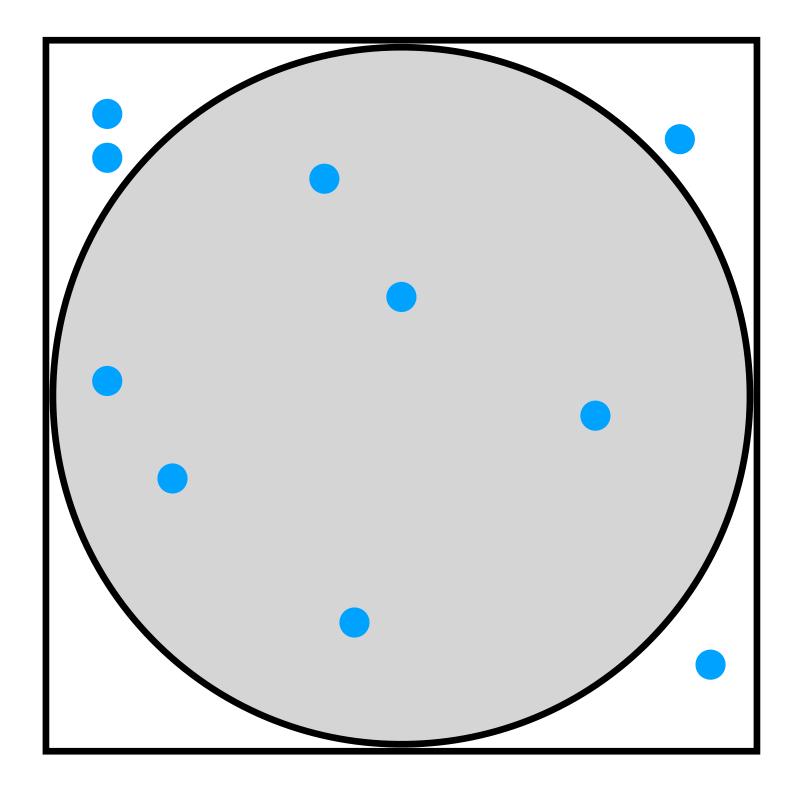


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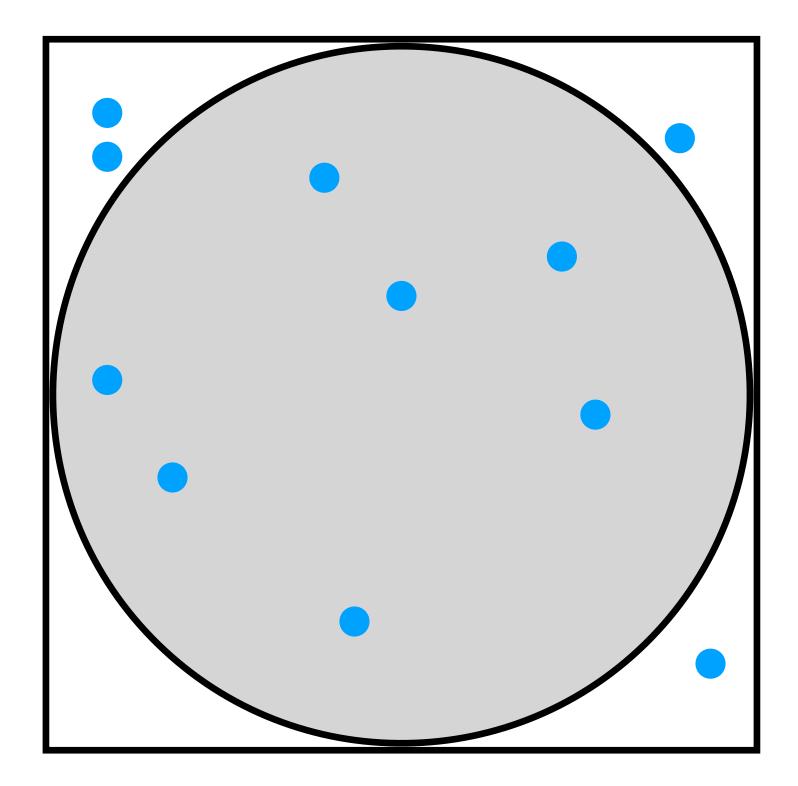


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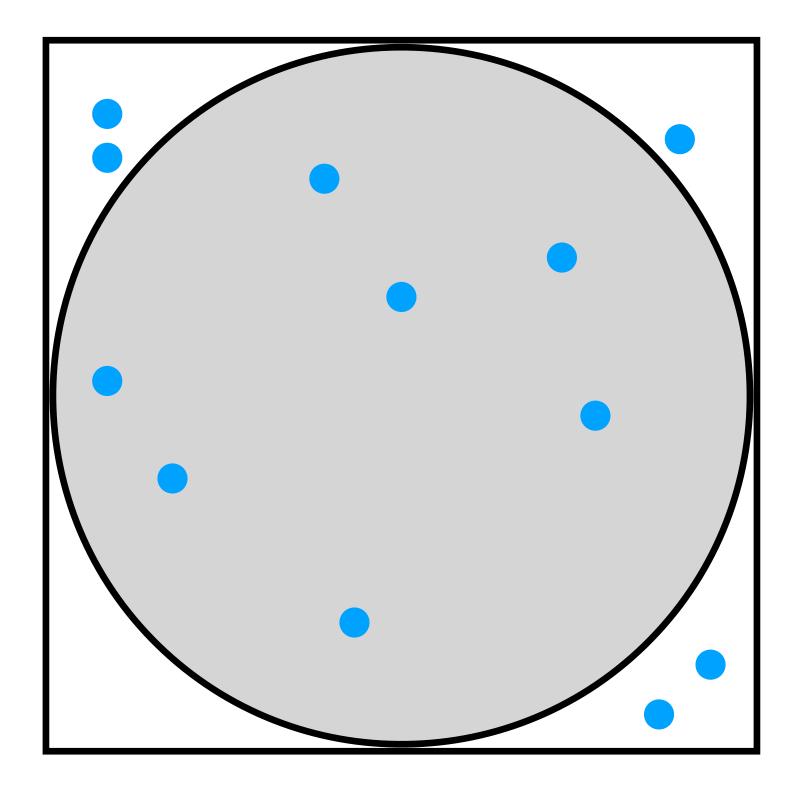


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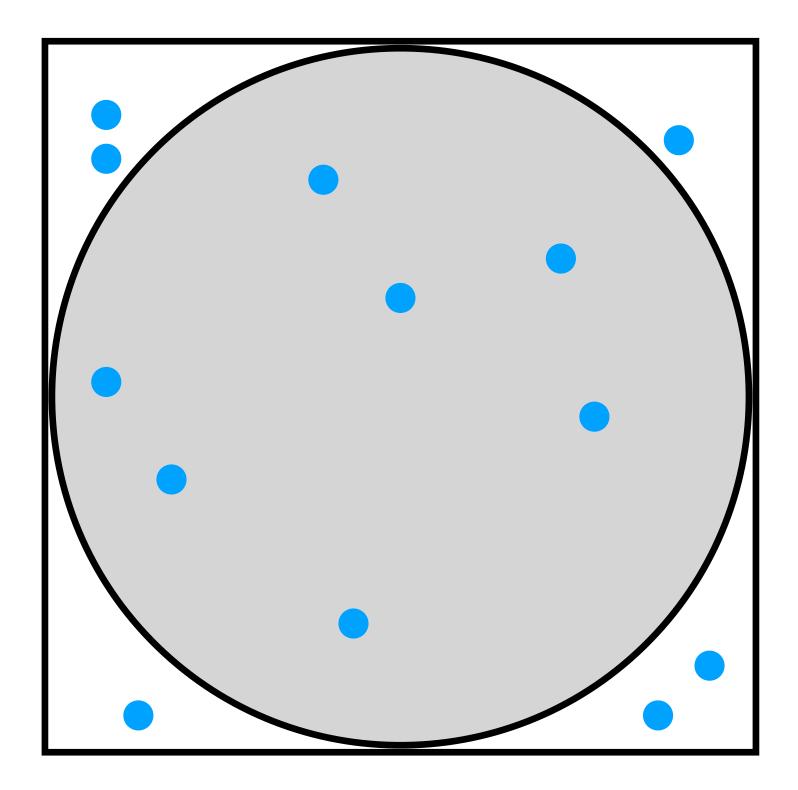


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Randomly simulate a process

P(desired outcome) = E[desired outcome]

simulations with desired outcome



$$\pi = 4 \frac{\text{area circle}}{\text{area square}} = 4 \frac{\text{inner}}{\text{total}}$$

Running example

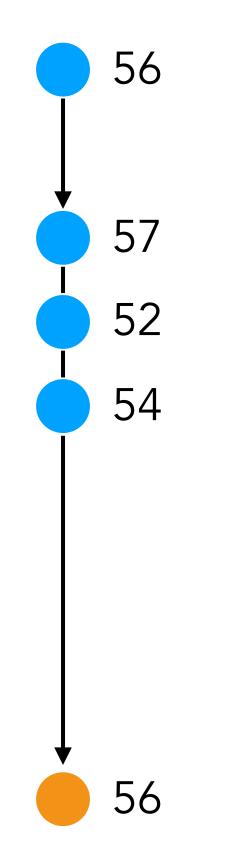
Measure a heat source in a factory with 3 different sensors

- heat ~ sample(Normal(56,10))
- sensor1 ~ sample(Normal(heat,3))
- sensor2 ~ sample(Normal(heat,5))
- sensor3 ~ sample(Normal(heat,5))
- observe(sensor2, Normal(43,2))
- return heat

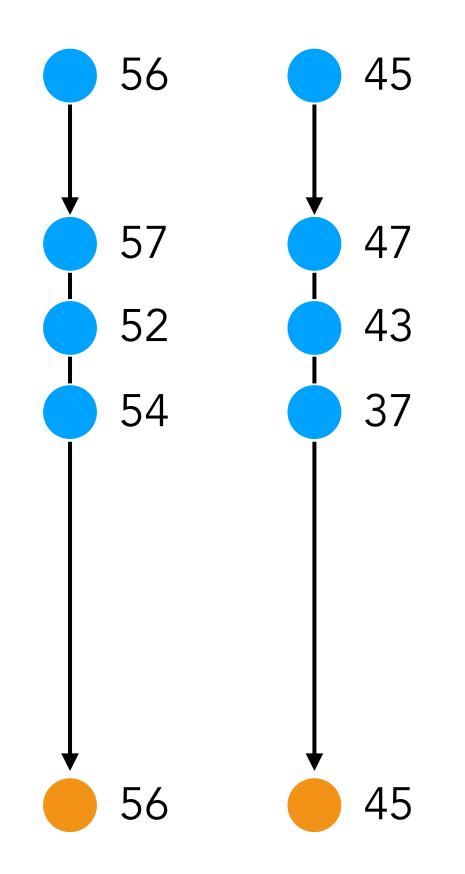
Probabilistic inference Importance sampling

- Execute the probabilistic program N times Treat the distribution over outcomes as the empirical distribution
- heat ~ sample(Normal(56,10))
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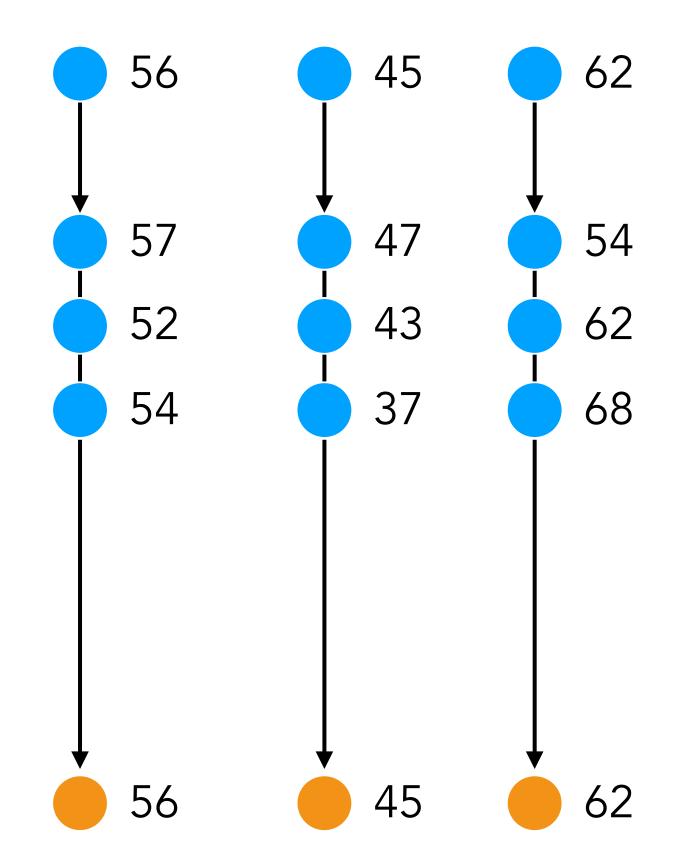
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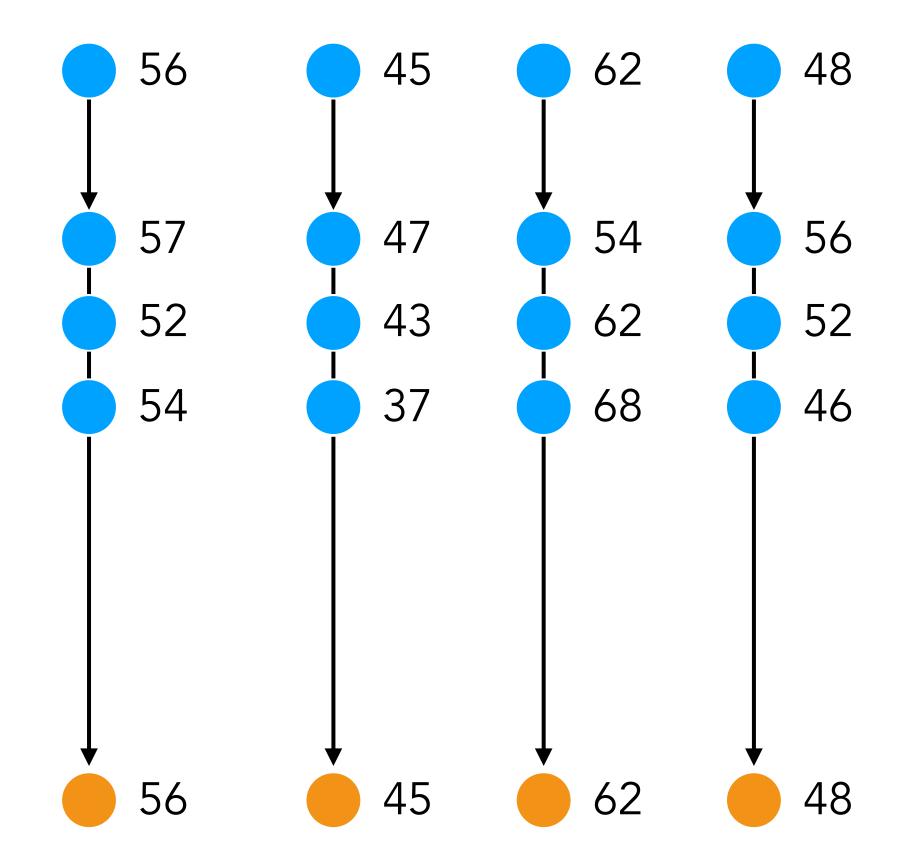
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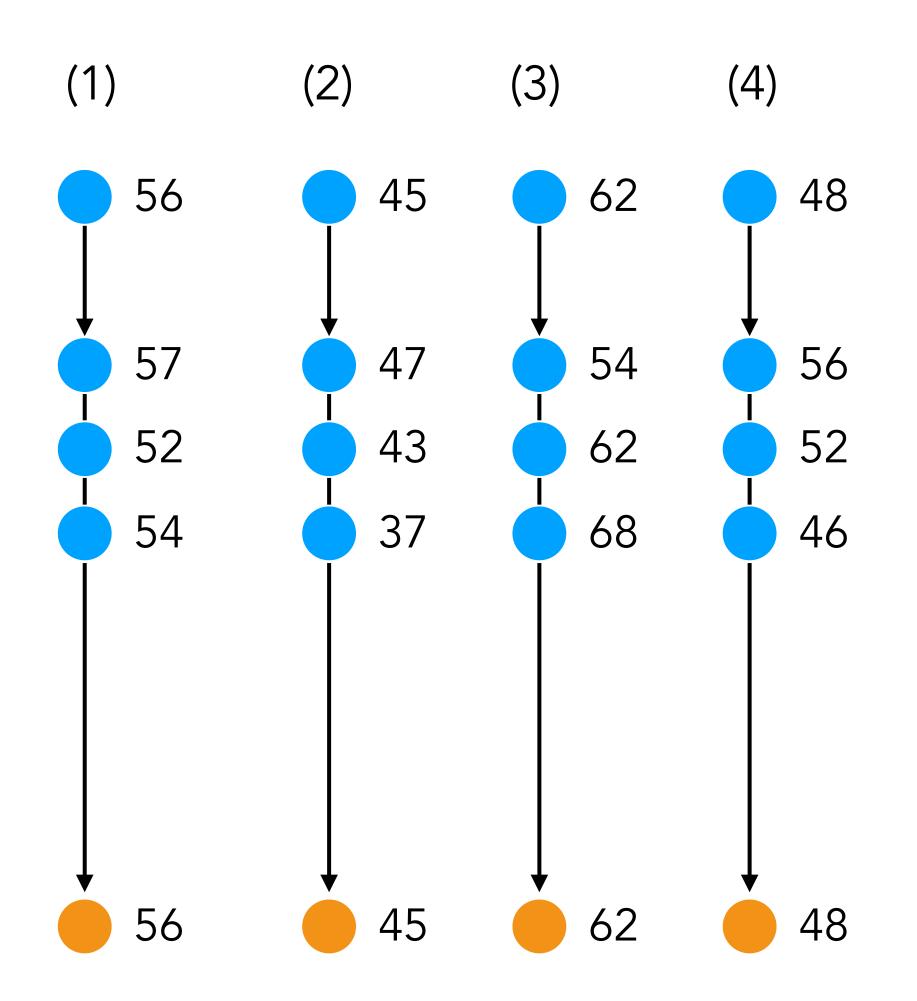
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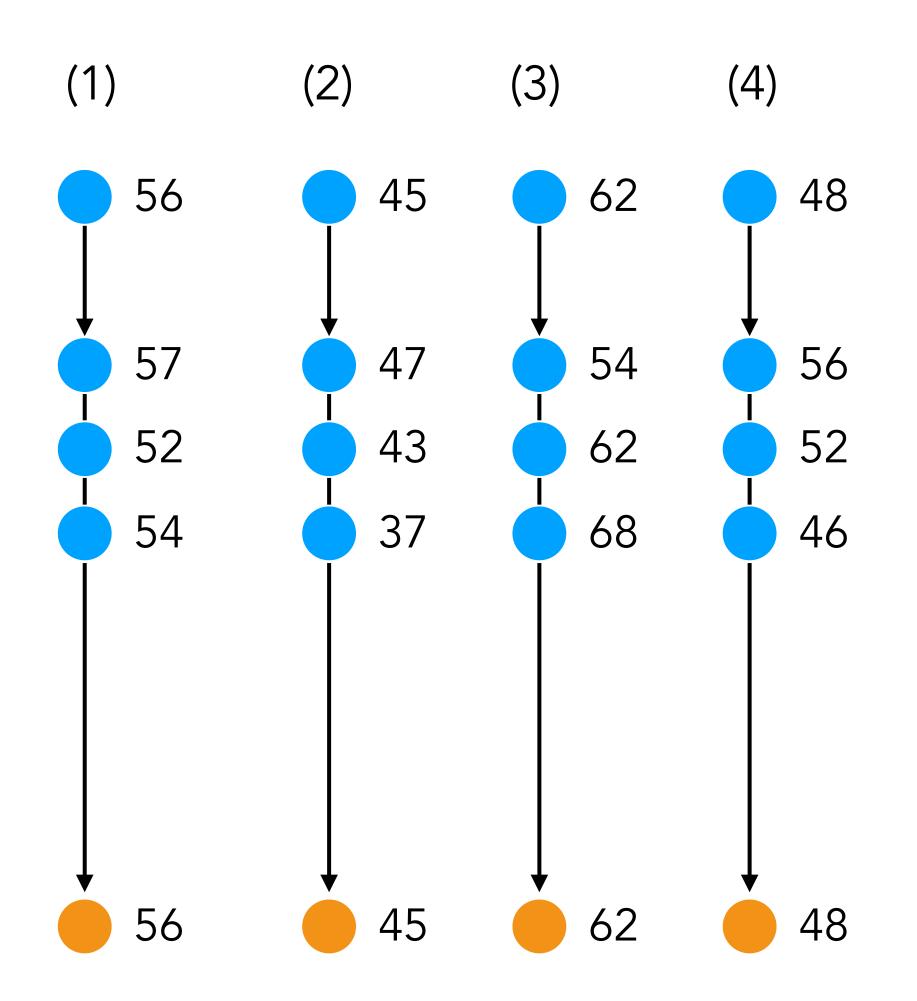


Samples obtained by executing a program are from p(heat) not $p(heat \mid sensor2 = 43)$!





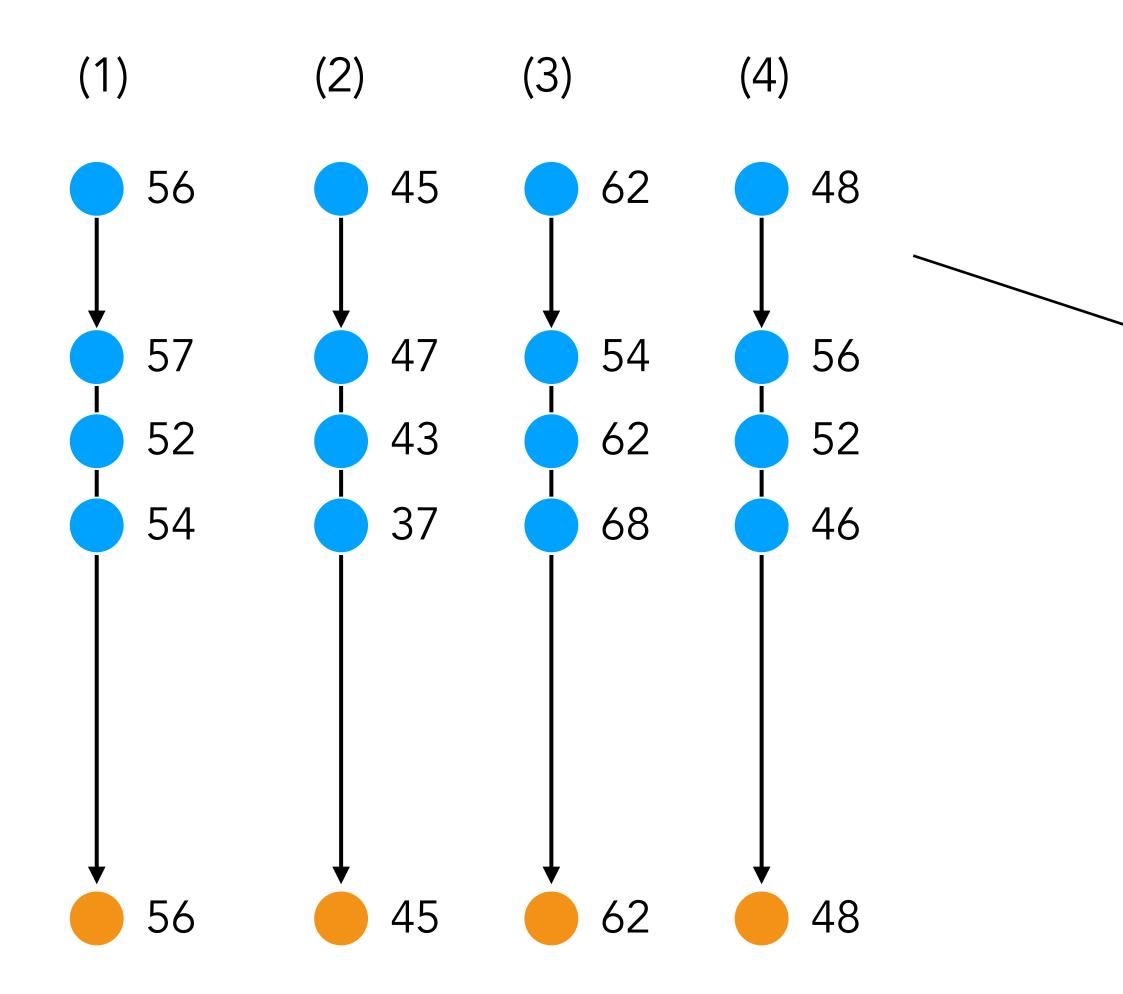
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We can fix this by weighting each execution proportionally to how much it agrees with observe(sensor2, Normal(43,2))



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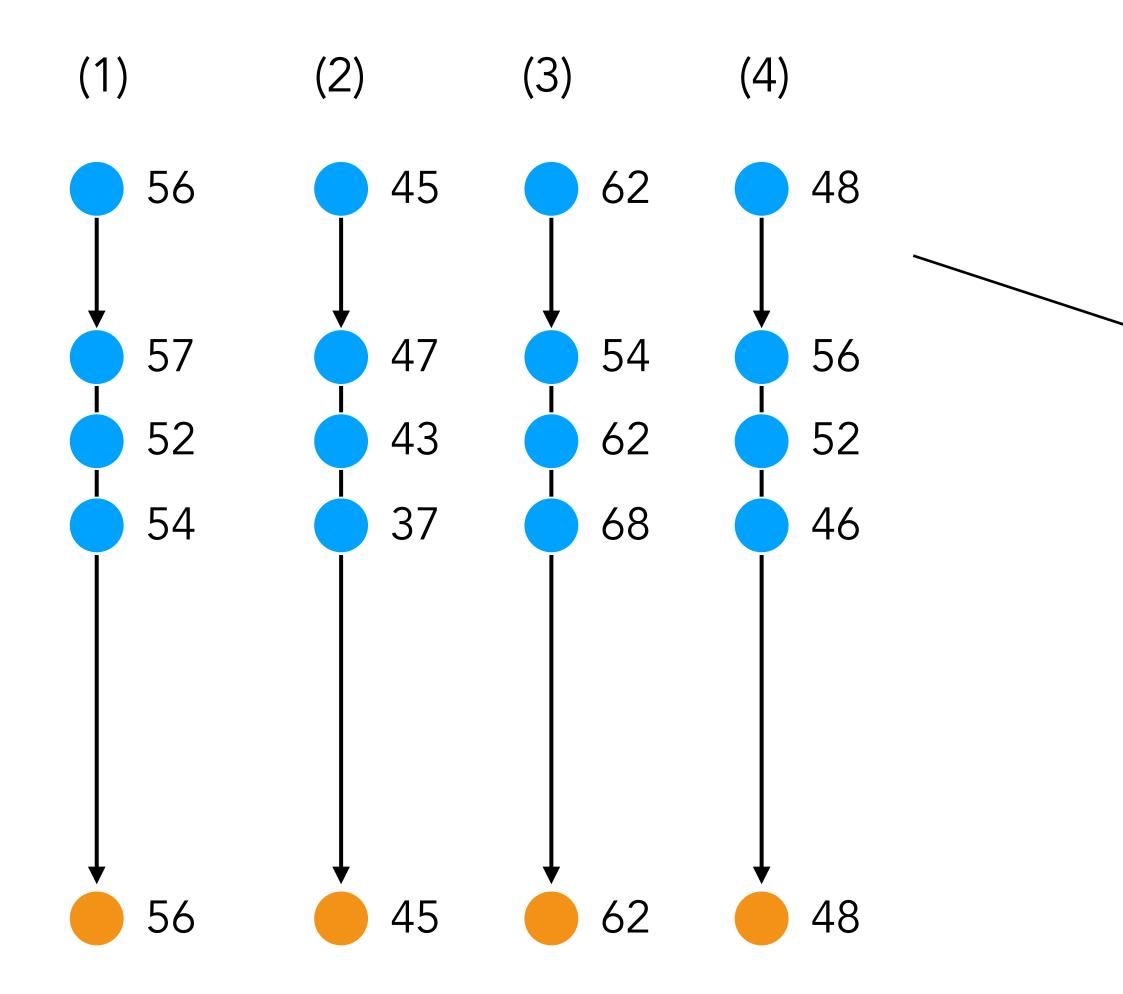


We can fix this by weighting each execution proportionally to how much it agrees with observe(sensor2, Normal(43,2))

> $heat^{4} = 48$ $sensor2^4 = 52$ $W^4 = p(sensor2 = 52; Normal(43,2))$



Samples obtained by executing a program are from p(heat) not $p(heat \mid sensor 2 = 43)$!

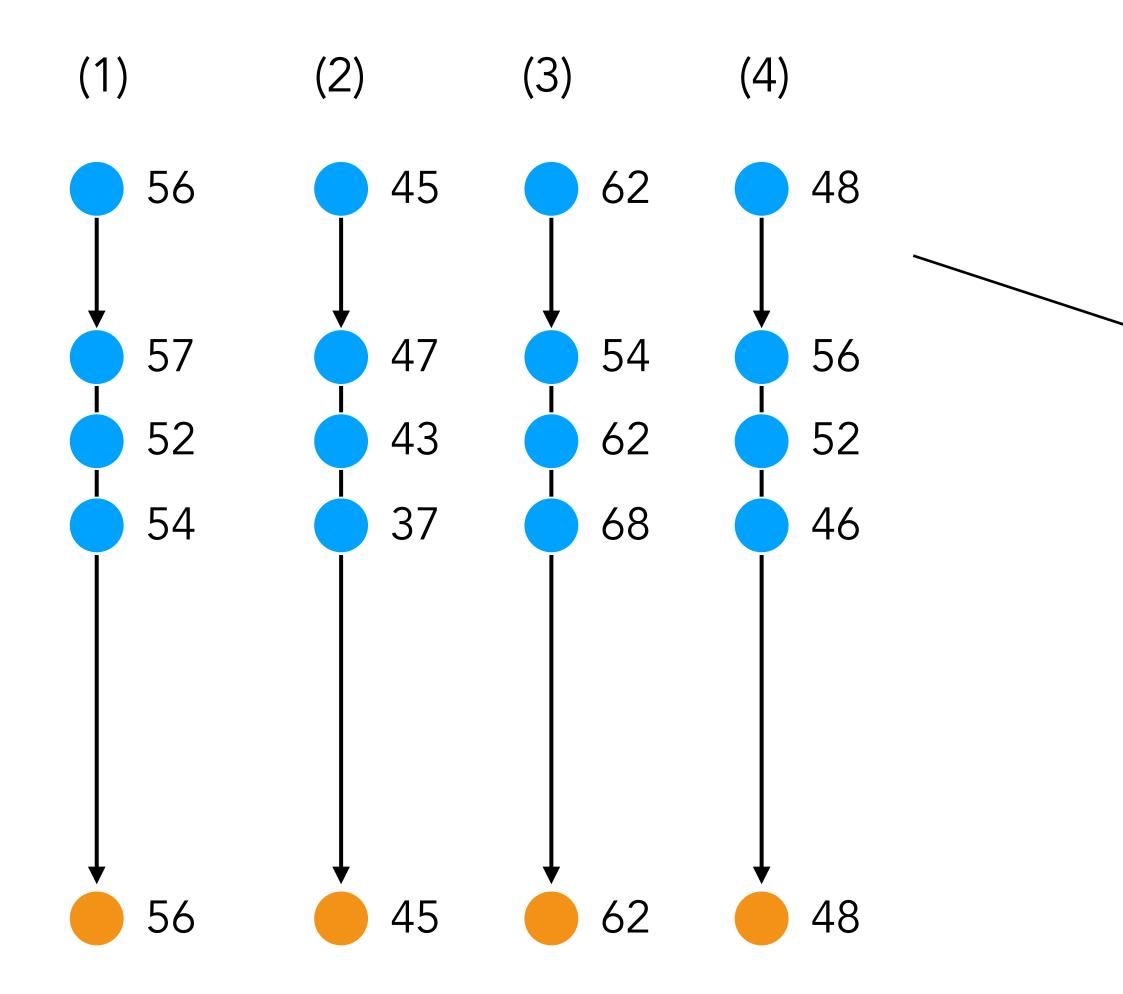


We can fix this by weighting each execution proportionally to how much it agrees with observe(sensor2, Normal(43,2))

 $W^3 = p(sensor2 = 62; Normal(43,2))$



Samples obtained by executing a program are from p(heat) not $p(heat \mid sensor 2 = 43)$!



We can fix this by weighting each execution proportionally to how much it agrees with observe(sensor2, Normal(43,2))

$$W^3 = p(sensor2 = 62; Normal(43,2))$$

Probability of any outcome *i* becomes:

 W^i $p(outcome^{l}) =$



Discrete and continuous expectations

$$\mathbb{E}[f] = \sum_{x} p(x) f$$

$$\mathbb{E}[f] = \int p(x)f$$

Conditional on another variable

$$\mathbb{E}_x[f|y] = \sum p$$

 \boldsymbol{x}

f(x)

(x) dx.

 $\partial(x|y)f(x)$

Sidestep sampling from the posterior $p(heat \mid sensor2 = 43)$ entirely, and draw from some proposal distribution q(heat) instead

Instead of computing an expectation with respect to p(heat | sensor2), We compute an expectation with respect to q(heat)

 $\mathbb{E}_{p(x|y)}[f(x)]$

Any distribution that is easy to sample from

$$\begin{aligned} &|=\int f(x)p(x|y)dx\\ &=\int f(x)p(x|y)\frac{q(x)}{q(x)}dx\\ &=\mathbb{E}_{q(x)}\left[f(x)\frac{p(x|y)}{q(x)}\right]\end{aligned}$$



We define an "importance weight"

Then with $x_i \sim q(x)$

$$\mathbb{E}_{p(x|y)}[f(x)] = \mathbb{E}_{q(x)}\left[f(x)W(x)\right] \approx \frac{1}{N}\sum_{i=1}^{N}f(x_i)W(x_i)$$

Expectations are now computed using weighted samples from q(x), instead of unweighted samples from p(x | y)

$$W(x) = \frac{p(x \mid y)}{q(x)}$$

One problem left: we cannot evaluate the weight just yet

$$W(x) = \frac{p(x \mid y)}{q(x)}$$

But we can evaluate it up to a constant

$$w(x) = \frac{p(x, y)}{q(x)}$$

Approximation

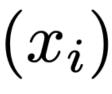
 $W(x_i) \approx \frac{w(x_i)}{\sum_{j=1}^N w(x_j)}$

We did all this to avoid calculating this term

 $\mathbb{E}_{p(x|y)}[f(x)] \approx \sum_{i=1}^{N} \frac{w(x_i)}{\sum_{j=1}^{N} w(x_j)} f(x_i)$







We already have a very simple proposal distribution we know how to sample from: the prior p(x)

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The algorithm then resembles the rejection sampling algorithm, except of sampling both the latest and the observed variables, we only sample the latent ones

Then, instead of a "hard" rejection step, we use the values of the latent variables and that data to assign "soft" weights to the sampled values



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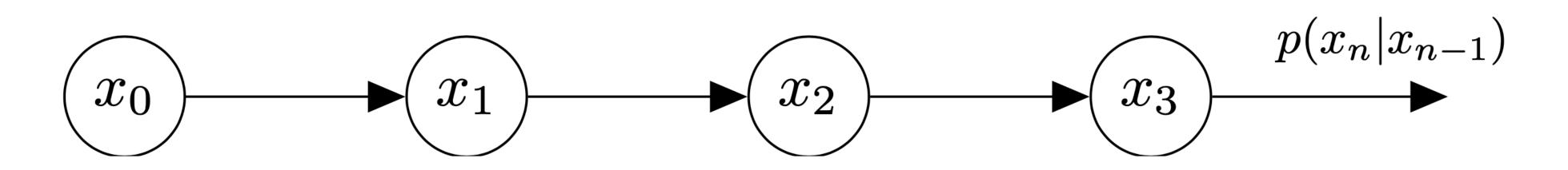
Importance sampling degrades poorly as the dimension of the latent variables increases, unless we have a very well-chosen proposal distribution q(x)

If the posterior distribution is 'peaky', we need a lot of luck to end up in the high-probability region

Probabilistic inference Metropolis-Hastings MCMC

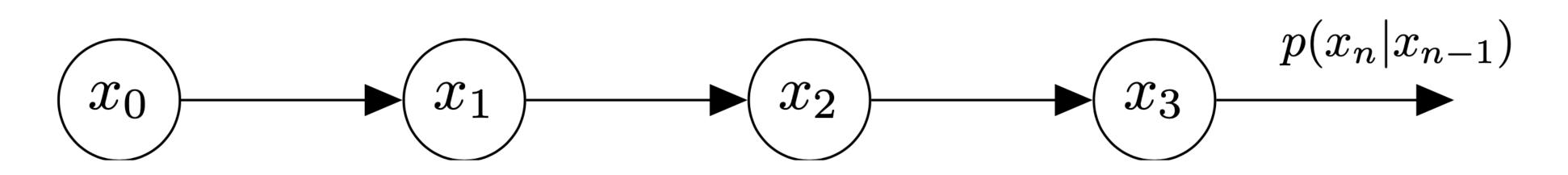
An alternative: Markov chain Monte Carlo methods draw samples from a target distribution by performing a biased random walk over the space of the latent variables x

The idea: create a Markov chain such that the sequence of states x_0, x_1, \ldots are samples from p(x | y)



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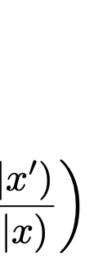
- One step = one sample (execution)

Use proposal distribution to make **local** changes to the latent variables (the trace). q(x'|x) then defines a conditional distribution over x' given a current value x

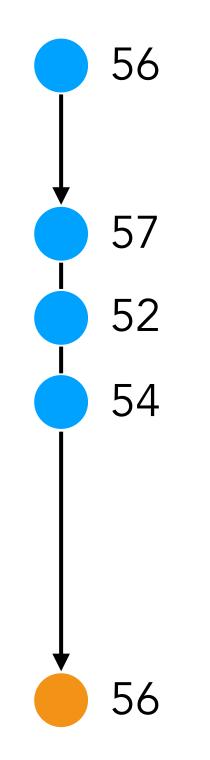
Do we keep the new trace?

$$A(x \to x') = \min\left(1, \frac{p(x', y)q(x)}{p(x, y)q(x')}\right)$$

Yes, with probability A



Generate the initial trace with e.g. IS

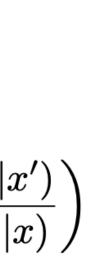


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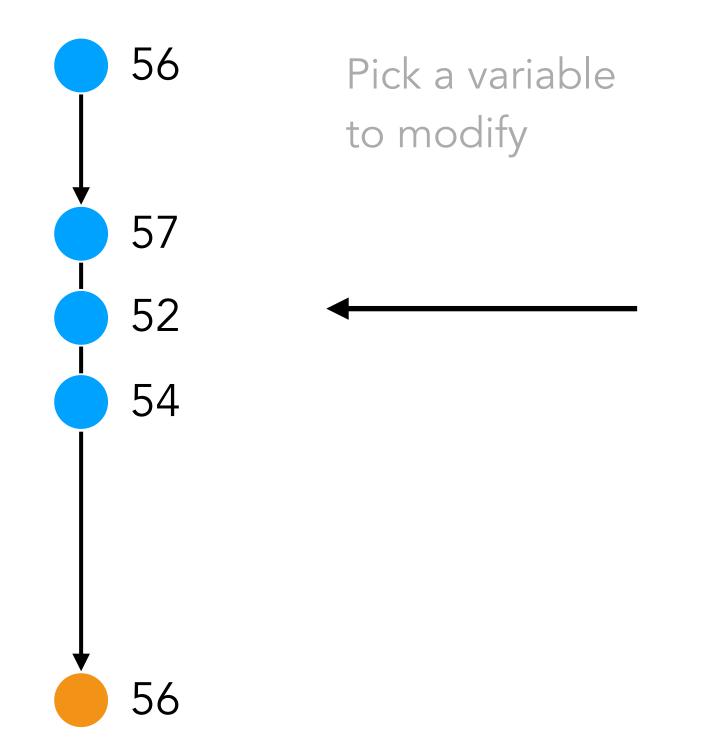
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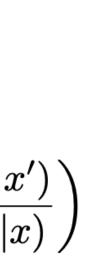


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Metropolis-Hastings

Generate the initial trace with e.g. IS

Pick a variable to modify

Modify the value by e.g. adding a small amount of noise

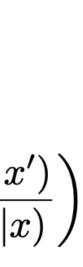
52 + Normal(0,2)

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Metropolis-Hastings

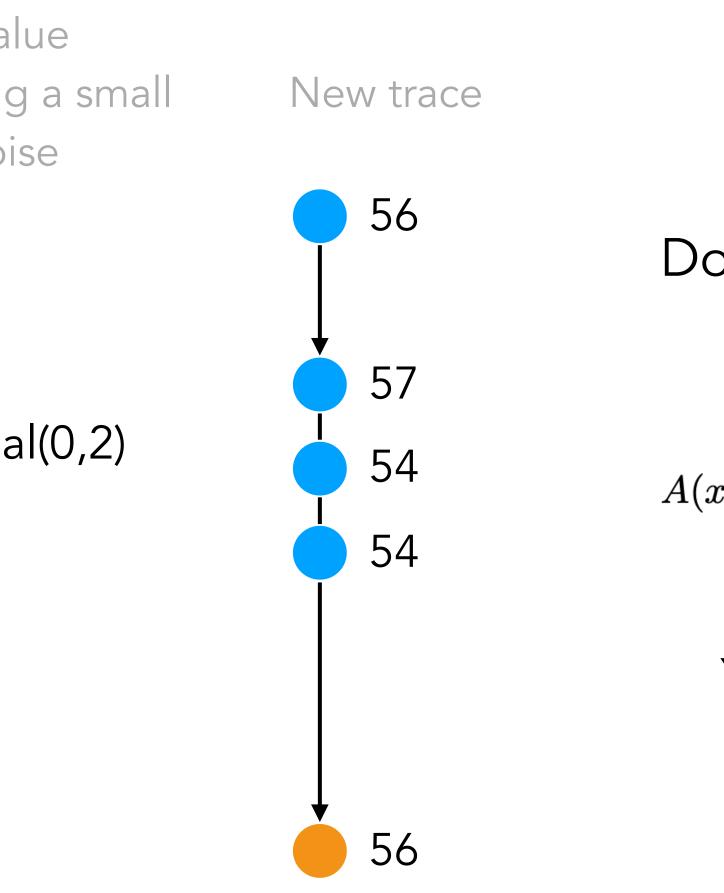
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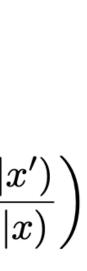
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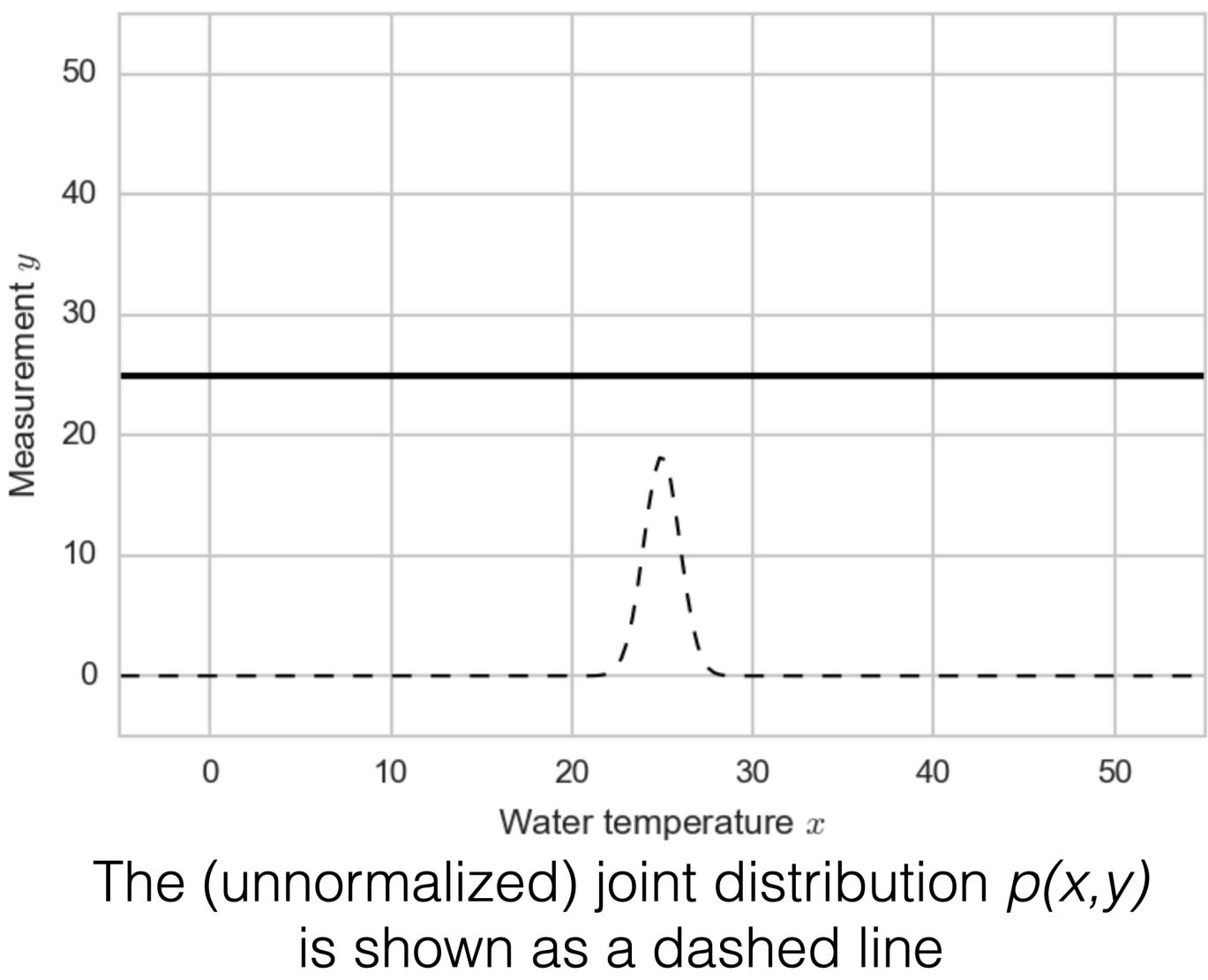


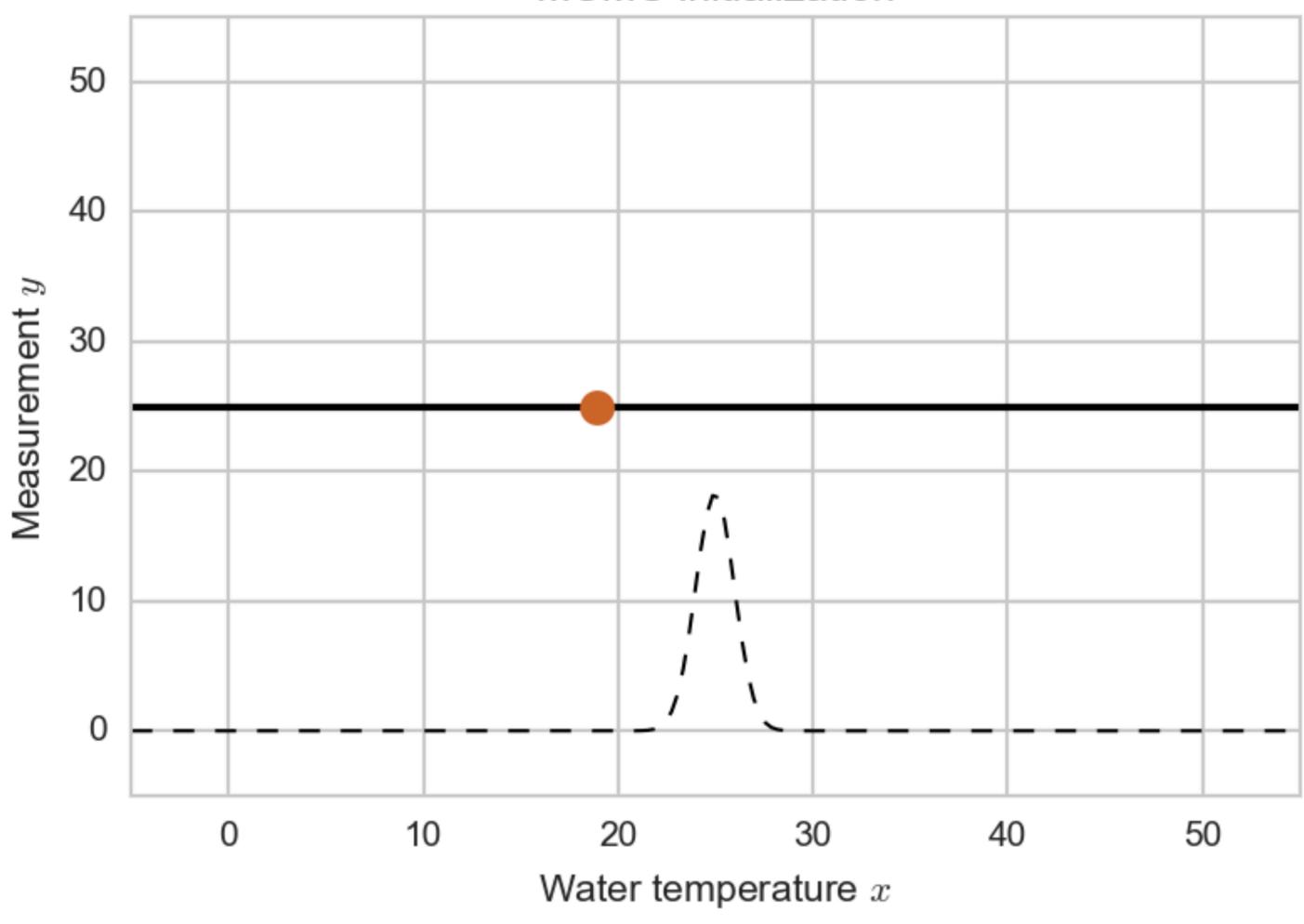
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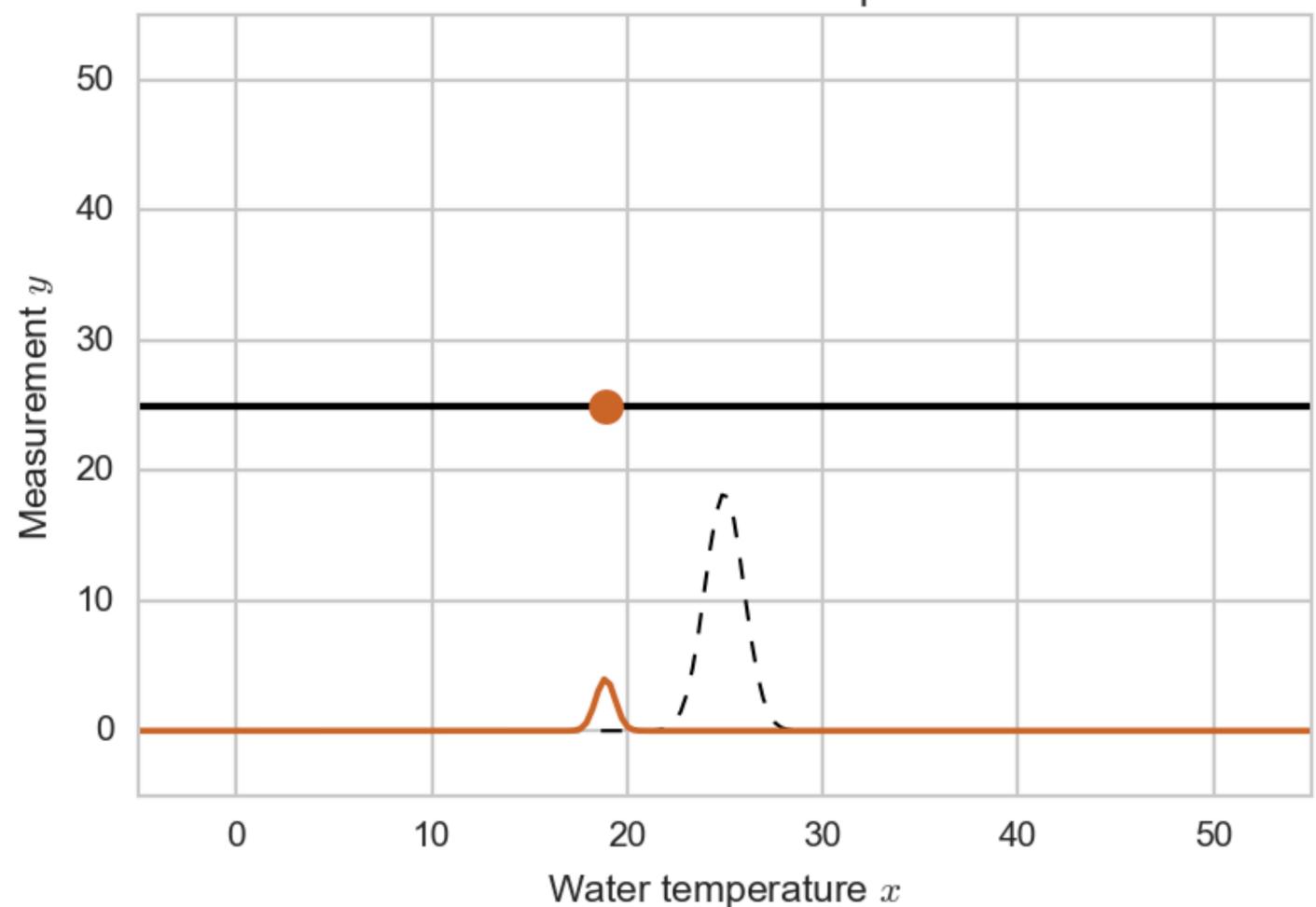






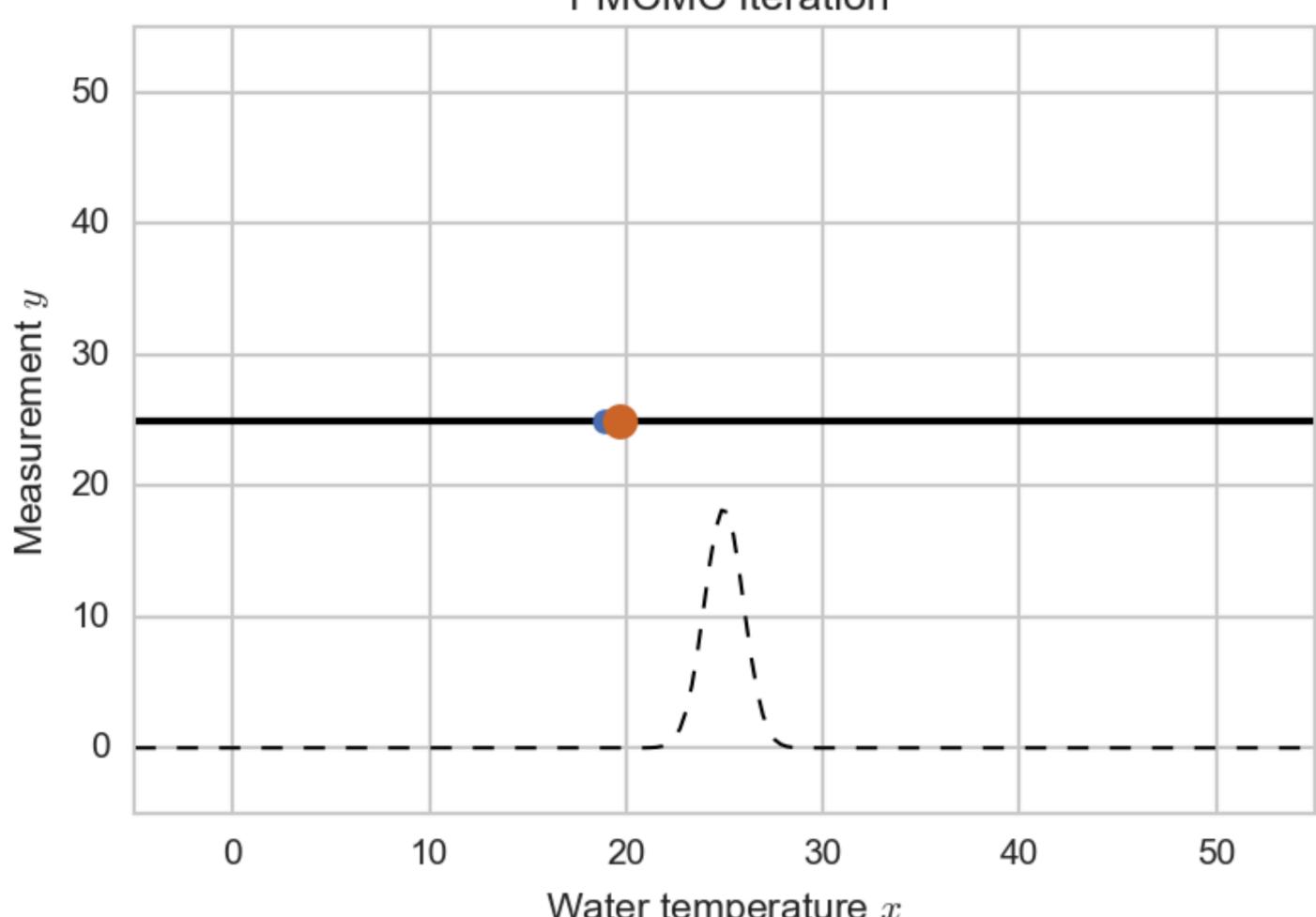
MCMC initialization

Initialize arbitrarily (e.g. with a sample from the prior)



Propose a local move on x from a transition distribution

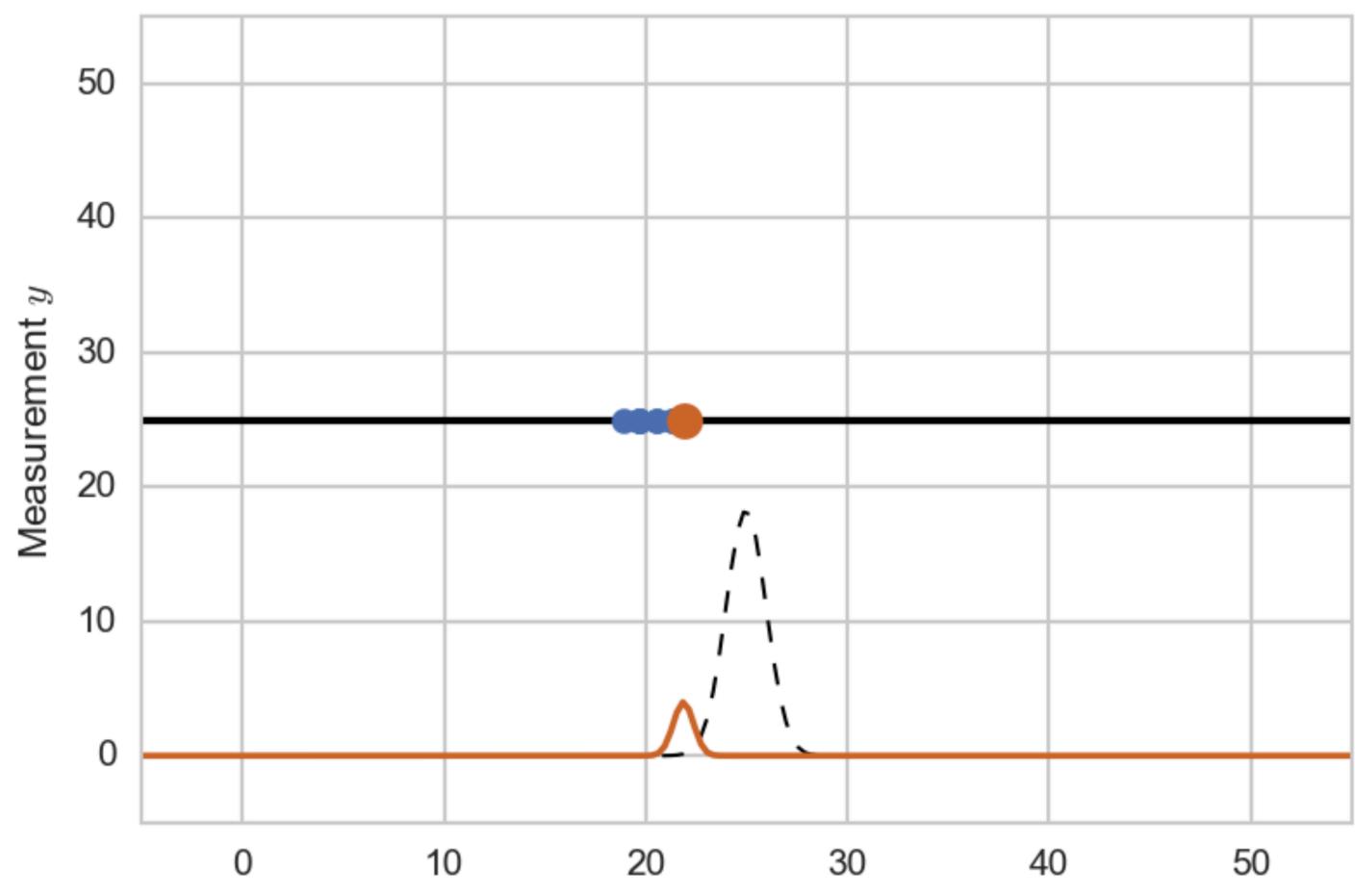
First MCMC step



Here, we proposed a point in a region of higher probability density, and accepted

1 MCMC iteration

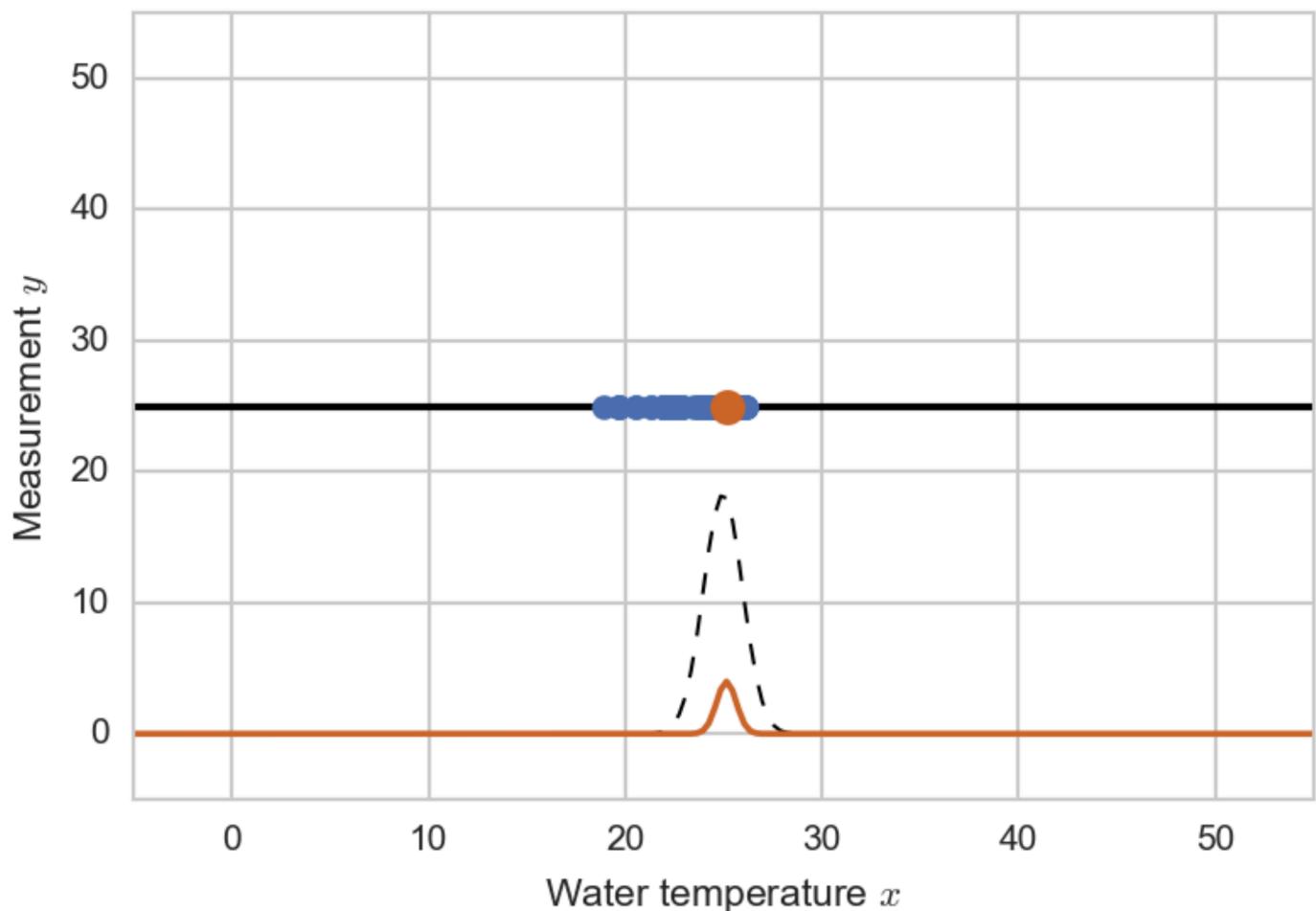
Water temperature x



Water temperature x

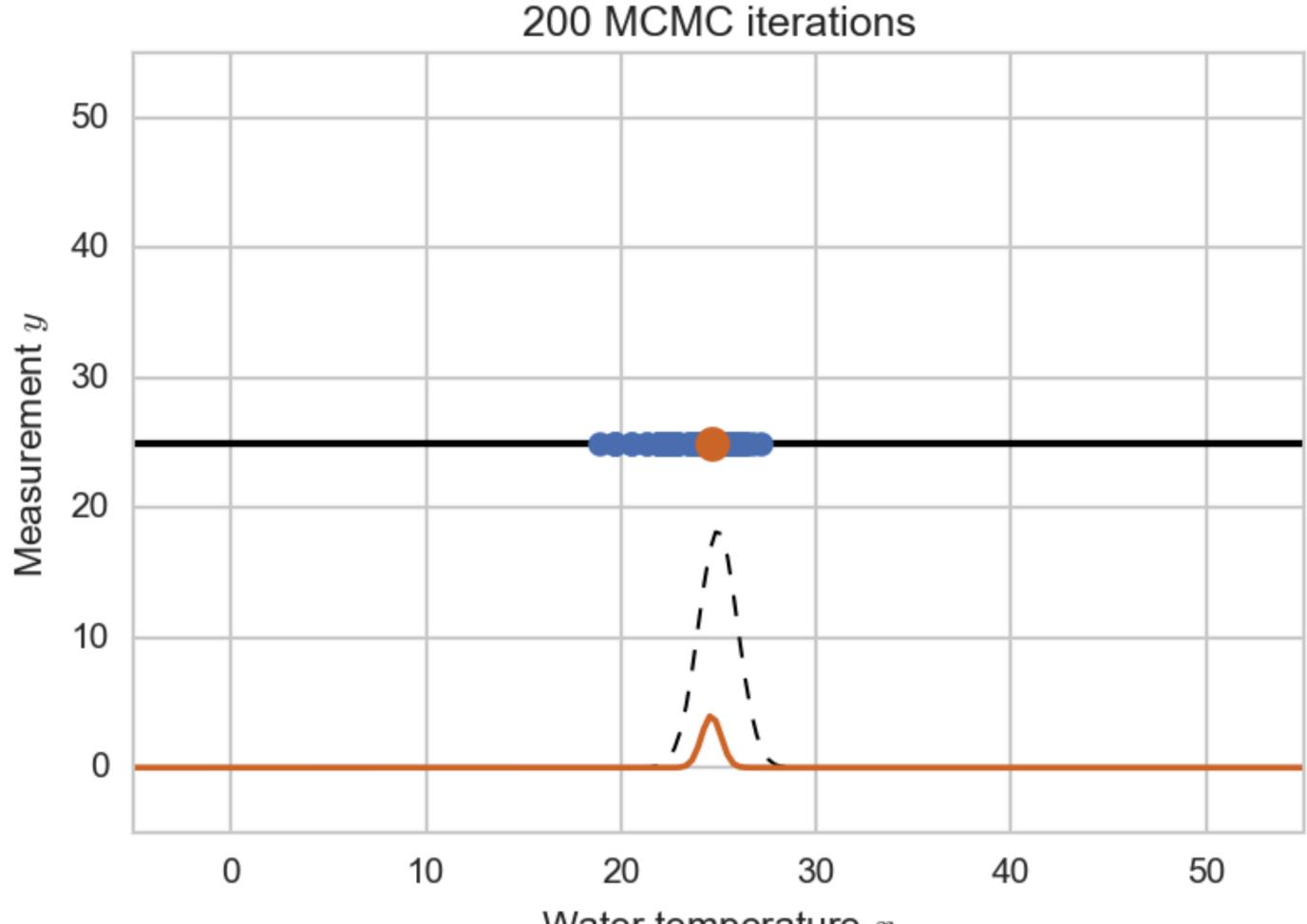
Continue: propose a local move, and accept or reject. At first, this will look like a stochastic search algorithm!

10 MCMC iterations



Once in a high-density region, it will explore the space

100 MCMC iterations



Once in a high-density region, it will explore the space

Water temperature x

Metropolis-Hastings MCMC: why can we re-weight?

The main technical requirement for MCMC is that the transition kernel leaves the posterior invariant

It is sufficient that the kernel satisfies the detailed balance criteria

q(X'|X, Y)p(X|Y) = q(X|X', Y)p(X'|Y)

Acceptance criterion ensures that!

If we sample $X \sim p(X | Y)$ and then generate a new sample $X' \sim q(X'|X, Y)$ from the transition kernel, X and X' come from the same distribution

We have to be able to go back to X from X'

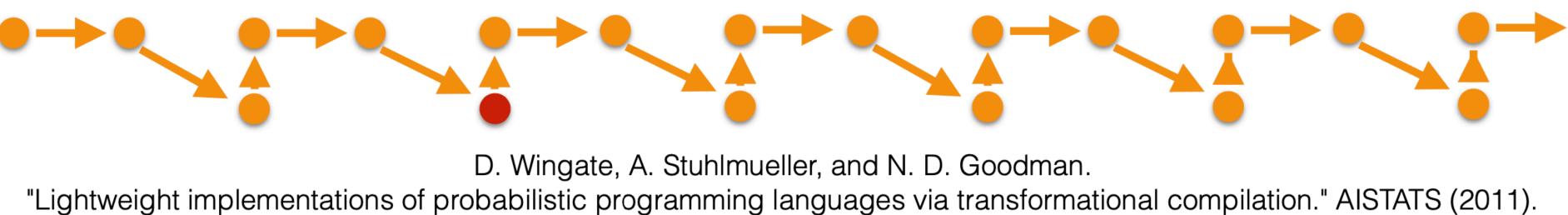
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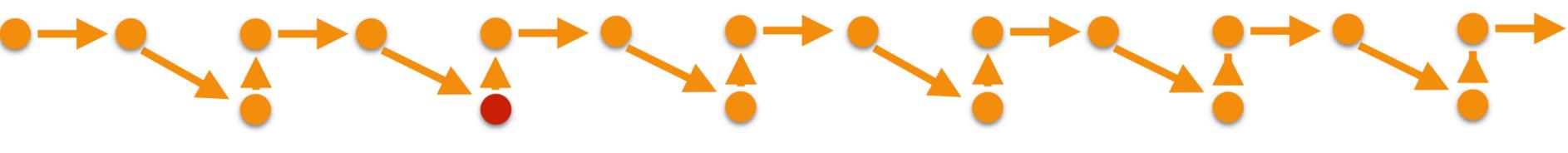
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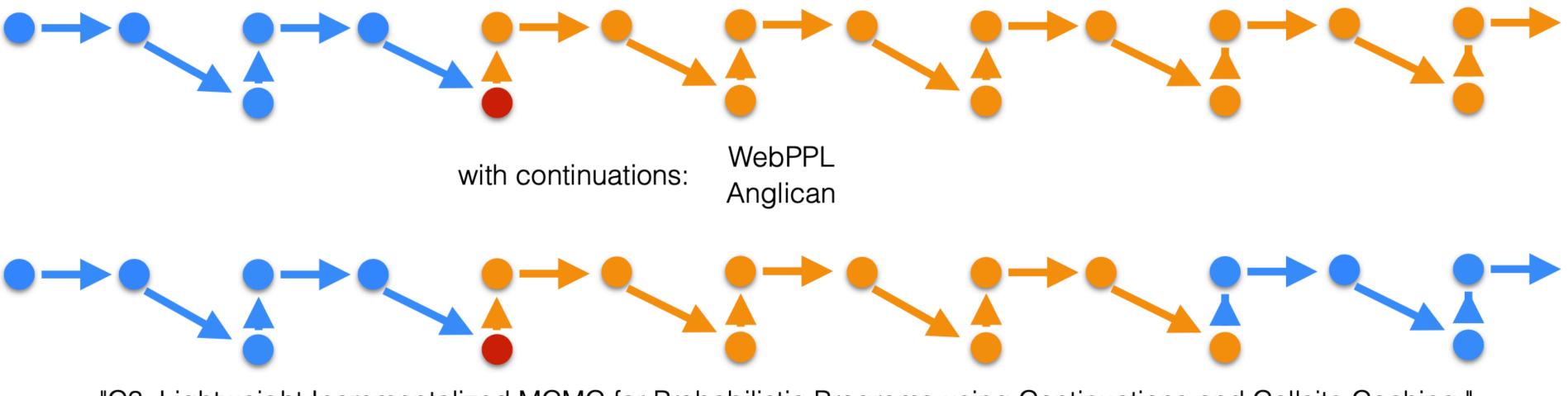
Metropolis-Hastings: computational efficiency

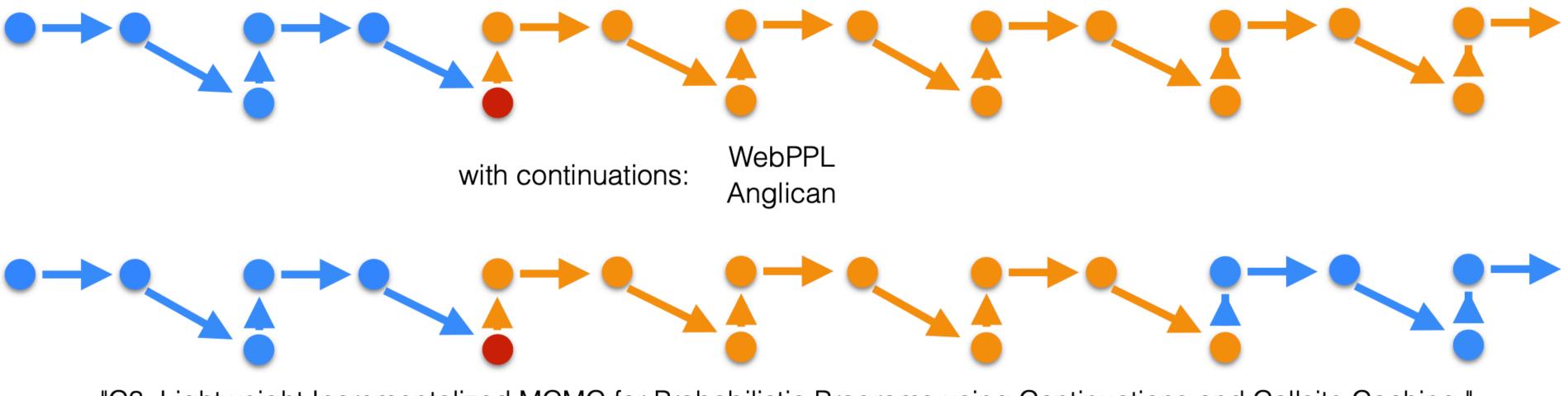


Metropolis-Hastings: computational efficiency



D. Wingate, A. Stuhlmueller, and N. D. Goodman. "Lightweight implementations of probabilistic programming languages via transformational compilation." AISTATS (2011).





"C3: Lightweight Incrementalized MCMC for Probabilistic Programs using Continuations and Callsite Caching." D. Ritchie, A. Stuhlmuller, and N. D. Goodman. arXiv:1509.02151 (2015).







Makes small changes to traces





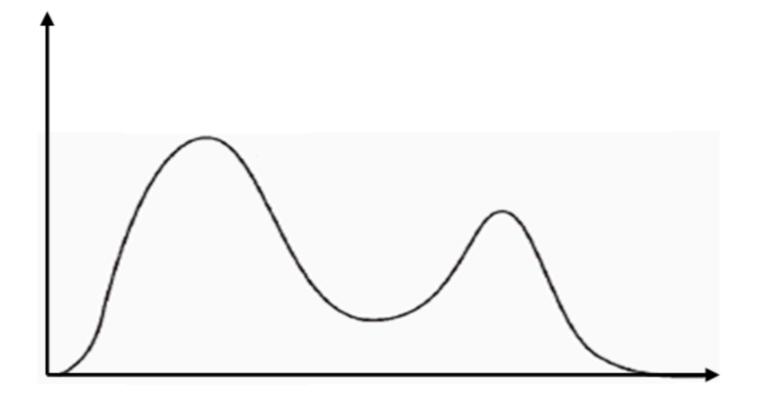
Makes small changes to traces Gradually goes to better traces





Makes small changes to traces Gradually goes to better traces





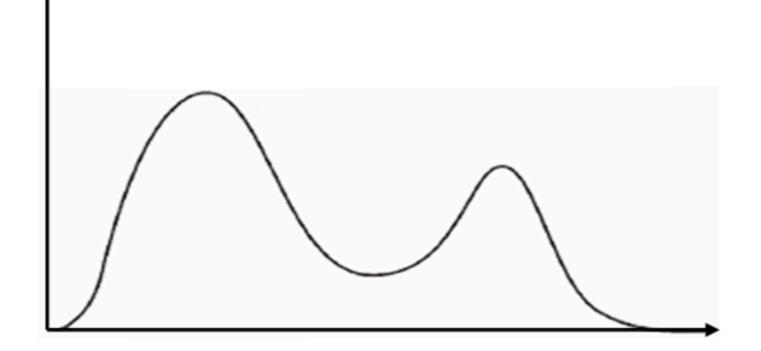


Makes small changes to traces Gradually goes to better traces



It might be difficult to capture a complex distribution in small steps

Especially when choices are correlated



Probabilistic inference Metropolis-Hastings MCMC

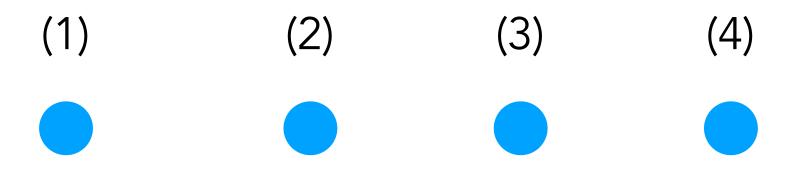
Importance sampling: makes all choices at once

Metropolis-Hastings: modify one choice at a time

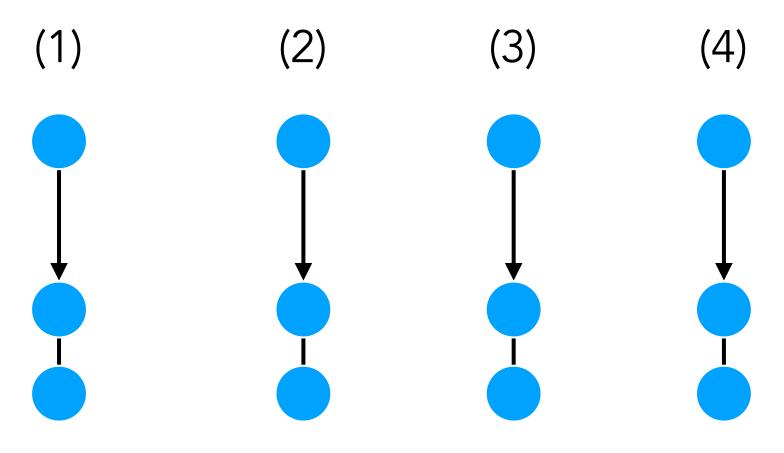
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Can we do better?

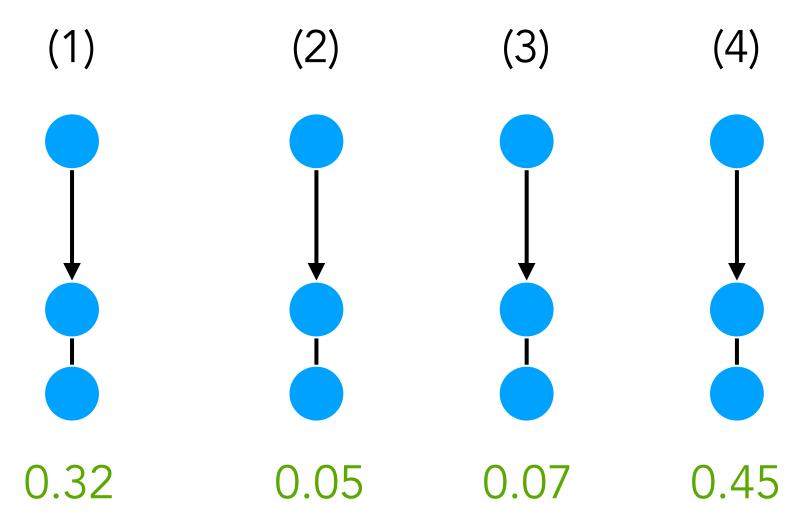


Initialise N traces/programs



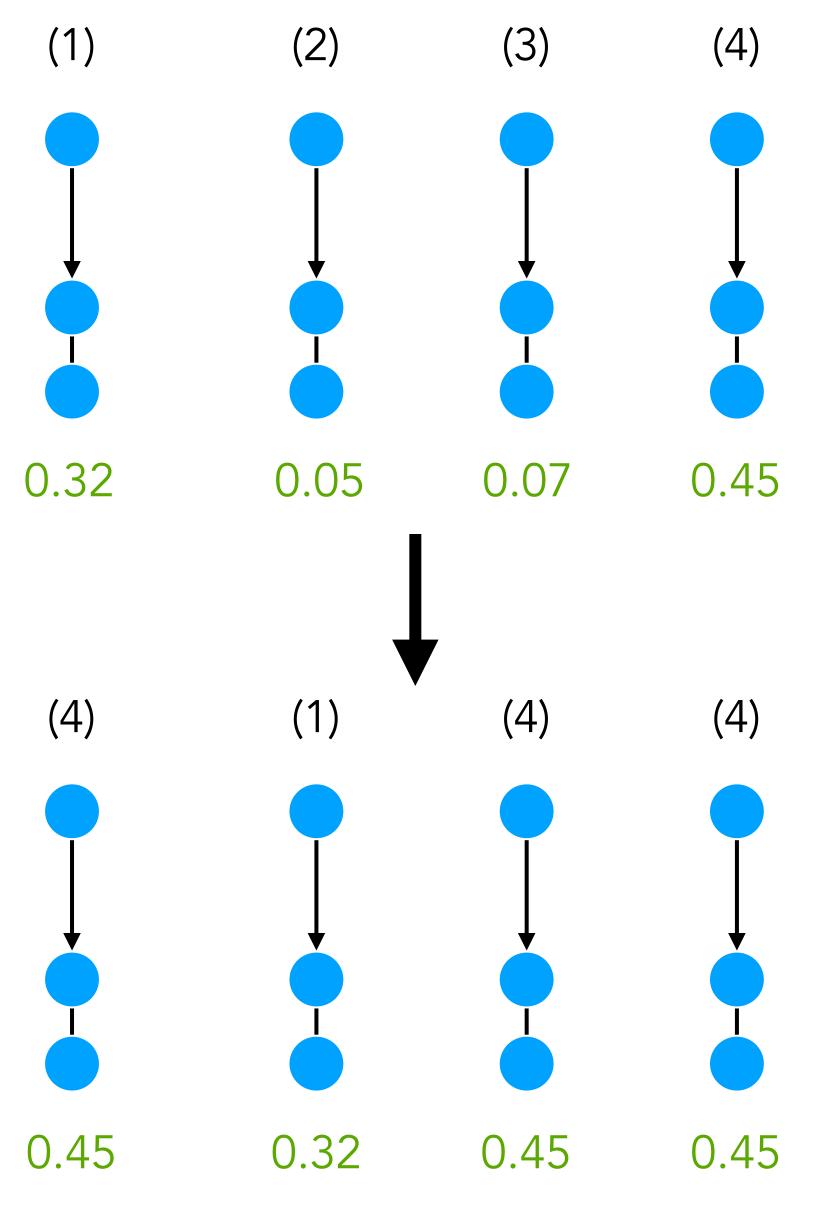
Initialise N traces/programs

Run them until the first observe statement



Initialise N traces/programs

Run them until the first observe statement And check how well they match the observations

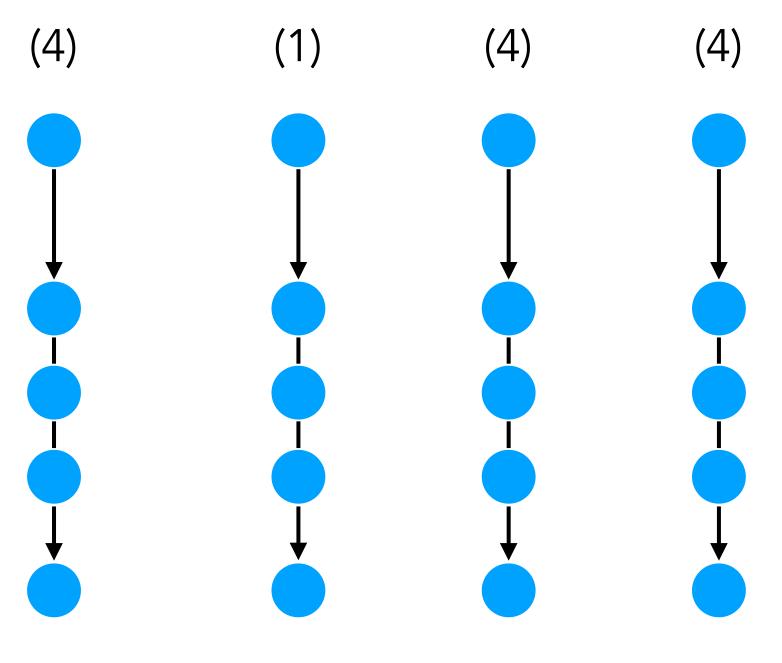


Initialise N traces/programs

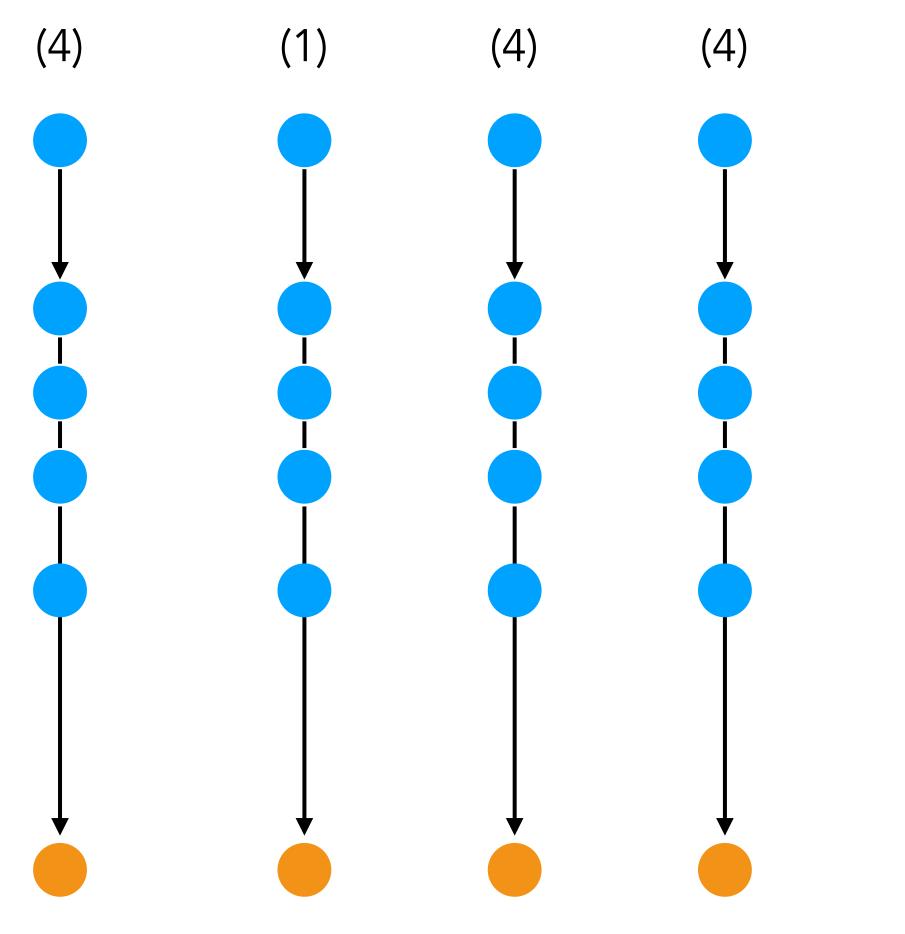
Run them until the first observe statement And check how well they match the observations

Then, resample **traces** with replacement proportional to how well they match observations





Run until the next observe and resample



Run until the next observe and resample

Continue until each program is finished

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Good mixture between IS and MH





Good mixture between IS and MH Often very good solution to complex programs





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Doesn't not support every programs



Good mixture between IS and MH Often very good solution to complex programs



Doesn't not support every programs What to do if all initial samples are bad?

i to complex program



Good mixture between IS and MH Often very good solution to complex programs



Doesn't not support every programs

What to do if all initial samples are bad?

 \rightarrow Particle filters with rejuvenation (perform Metropolis-Hastings) on the traces before sampling)





Calculating p(x, y) exactly is not possible for non-toy problems

We have to rely on Monte Carlo approximations

Inference procedures need to be able to handle any kind of program Importance sampling

Metropolis-Hastings MCMC

Particle filtering